Pricing and Hedging GLWB in the Heston and in the Black-Scholes with Stochastic Interest Rate Models*

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Abstract

Valuing Guaranteed Lifelong Withdrawal Benefit (GLWB) has attracted significant attention from both the academic field and real world financial markets. As remarked by Forsyth and Vetzal [9] the Black and Scholes framework seems to be inappropriate for such long maturity products. They propose to use a regime switching model. Alternatively, we propose here to use a stochastic volatility model (Heston model) and a Black Scholes model with stochastic interest rate (Hull White model). For this purpose we present four numerical methods for pricing GLWB variables annuities: a hybrid tree-finite difference method and a hybrid Monte Carlo method, an ADI finite difference scheme, and a standard Monte Carlo method. These methods are used to determine the no-arbitrage fee for the most popular versions of the GLWB contract, and to calculate the Greeks used in hedging. Both constant withdrawal and optimal withdrawal (including lapsation) strategies are considered. Numerical results are presented which demonstrate the sensitivity of the no-arbitrage fee to economic, contractual and longevity assumptions.

Keywords: Variable Annuities, stochastic volatility, stochastic interest rate, optimal withdrawal.

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1 Introduction

In 2008 following the subprime crisis, financial markets have suffered the upheavals that have affected the entire world economy. Since then, these markets were extremely volatile: this situation could last a while, and perhaps become the new standard. After many failures, the gap between the different interest rate apply to different transmitters has become larger and larger, and a discussion on the identification of the risk-free rate is opened. The ECB and the Fed's rate gradually declined, while rate on sovereign debt increased gradually.

For customers, it is difficult to balance risk and return. In this context, clients seek protection for their savings, and the ability to take advantage of the positive changes in the market. With regard social problematic, following the increase in life expectancy, annuities for retirement dropped.

The mission of insurance companies is to answer the request for protection and compensation of its customers. The solution is to provide the customer an investment account and cover its value with guarantees. These products are called Variable Annuities. In the words of François Robinet, CEO of AXA Life Invest, "These products, unit of account guaranteed will become a solution to solve the long-term investment problems with security, prepare for retirement".

In this article, we consider a Guaranteed Lifelong Withdrawal Benefit (GLWB) annuity. We restrict our attention to a simplified form of a GLWB which is initiated by making a lump sum payment to an insurance company. This lump sum is then invested in risky assets, usually a mutual fund. The benefit base, or guarantee account balance, is initially set to the amount of the lump sum payment. The holder of the contract is entitled to withdraw a fixed fraction of the benefit base for life, even if the actual investment in the risky asset declines to zero. Upon the death of the contract holder, his (her) estate receives the remaining amount in the risky asset account. Typically, these contracts have ratchet provisions (step-ups), that periodically increase the benefit base if the risky asset investment has increased to a value larger than the guarantee account value, and roll up provisions, that periodically increase the benefit base according to a deterministic function. In addition, the benefit base may also be increased if the contract holder does not withdraw in a given year (bonus). Finally, the contract holder may withdraw more than the contractually specified amount, including complete surrender of the contract, upon payment of a penalty. Complete surrender here means that the contract holder withdraws the entire amount remaining in the investment account, and the contract terminates. In most cases, this penalty for full or partial surrender declines to zero after five to seven years.

The hedging costs for this guarantee are offset by deducting a proportional fee from the risky asset account. From an insurance point of view, these products are treated as financial ones: the products are hedged as if they were pure financial products, and the mortality risk is hedged using the law of large numbers. Therefore, it is very important for insurance companies to be able to price quickly these products. Moreover these products have long maturities that could last almost 60 years. The Black-Scholes model, with constant interest rate and volatility seems to be unsuitable for those products: that's why we present our pricing methods in two frameworks, modeling stochastic volatility (Heston model [11]) and stochastic interest rate (Hull-White model [14]).

There have been several recent articles on pricing GLWBs. In particular, we would remember the Forsyth and Vetzal's work [9]: they used a PDE approach in a multi regimes model to price GLWBs contracts. This approach proved to be very fast and accurate, and we used it as a reference for our work. Concerning the use of stochastic volatility, Kling et al. [15] used a Monte Carlo approach to price products. We have made reference also to Bacinello et al. [3]: variable annuities (including GLWBs) are priced using a Monte Carlo approach. The policy holder (hereinafter, we will abbreviate it with PH) behavior is assumed to be semi-Static, i.e. the holder withdraws at the contract rate or surrenders the contract.

In this paper, we price GLWBs guarantees, and we find the no-arbitrage fee, in the Heston model and the Black-Scholes with stochastic interest rate model (BS HW model). First, we treat a static withdrawal strategy: the PH withdraws at the contract rate. Then, taking the point of view of the worst case for the hedger, we price the guarantees assuming that the contract holder follows an optimal withdrawal strategy. We also used these methods to calculate the Greeks for hedging and risk management. Moreover we performed a mortality shock useful in risk management framework. For this purpose we present four numerical methods:

a hybrid tree-finite difference method and a hybrid Monte Carlo method (both introduced by Briani et al. [4]) an ADI finite difference scheme (Haentjens and Hout [13]), and a Standard Monte Carlo method with Longstaff-Schwartz least squares regression (Longstaff and Schwartz [16]).

We use the term *no-arbitrage fee* in the sense that this is the fee which is required to maintain a replicating portfolio. A description of the replicating portfolio for these types of guarantees is given in Chen et al. [8] and Belanger et al. [6].

The main results of this paper are the following ones:

- We formulate the determination of the no-arbitrage fee (i.e. the cost of maintaining a replicating hedging portfolio) in the Heston model and in the BS HW model using different pricing methods;
- We present the effects of stochastic volatility and stochastic interest rate on pricing and Greeks calculation, and the sensitivity of the GLWB fee to various modeling parameters;
- We use different numerical methods to price the GLWB contract;
- We present numerical examples which show the convergence of these methods.

The paper is organized as follows: in Section 2, we describe the main features of the contract such as mortality, withdrawals, and ratchets. In Section 3, we provide a brief review of the stochastic models used afterward. In Section 4, we present the numerical methods, and how to implement them to solve the GLWB contract pricing problem. In Section 5 we perform tests in order to show their behavior and we study the sensitivity of the no-arbitrage fee to economic, contractual and longevity assumptions. Finally, in Section 6, we present the conclusions.

2 The GLWB Contract

In the following, we will refer to the contract described in the paper of Forsyth [9], with some variations useful to compare our results with other works. We make a brief summary of the main features of the contract.

2.1 Mortality

We price the products in a risk-neutral measure, therefore in the following we assume that mortality risk is diversifiable (Milevsky and Salisbury, [17]). When this assumption is not justified, then the risk-neutral value of the contract can be adjusted using an actuarial premium principle (Gaillardetz and Lakhmiri, [10]). Hereinafter, the time variable will be denoted by the letter t, and we assume that the contract starts at t = 0.

First we suppose that no PH can live longer than a given age. This age will be denoted by τ (usually $\tau = 122$). The age of the PH at the beginning will be denoted by a_0 (usually $a_0 = 65$). So, the maturity of the contract is $T = \tau - a_0$ (usually T = 57): when the time variable t reaches T all PHs are died, and the contract is worth zero.

The effects of the mortality on the contract are described using two functions:

- $\mathcal{M}:[0,T]\to\mathbb{R}$ is the probability density that describes the random variable M associated to the death year of the PH. The fraction of the original owners who die in [t,t+dt] is equal to $\mathcal{M}(t) dt$.
- $\mathcal{R}: [0,T] \to \mathbb{R}$ is the fraction of the original owners who are still alive at time t

$$\mathcal{R}\left(t\right) = 1 - \int_{0}^{t} \mathcal{M}\left(s\right) ds.$$

We remark that $\mathcal{R}(0) = 1$ and $\mathcal{R}(T) = 0$. For seek of simplicity, we assume \mathcal{M} to be constant between contract's anniversaries: if $t \in [k, k+1[$, $k \in \mathbb{N}$ then $\mathcal{M}(t) = \mathcal{M}(k)$.

2.2 Contract State Parameters

At time t = 0 the policy holder pays with lump sum the gross premium GP, to the insurance company. This may be reduced by some initial fees, giving a net premium P. The premium P is invested in a fund, whose price is denoted by the variable S_t . The state parameters of the contract are:

• Account value: A_t , $A_0 = P$.

• Base benefit: B_t , $B_0 = GP$.

Both these two variables are initially set equal to the gross premium or to the premium.

We suppose that the acquisition charges are equal to GP - P aren't used for hedging purposes, but only to cover entry costs for management control. We suppose that there is a set of discrete times t_i , which we term *event times*. At these times, withdrawals, ratchets, and bonuses may occur. Normally, event times are annually or quarterly. We first consider the evolution of the value of the guarantee excluding these event times t_i .

The value of the contract at time t is denoted by $V(A_t, B_t, t)$.

2.3 Evolution of the Contract between Event Times.

Let $t \in]t_i, t_{i+1}[\subseteq [0, T]]$. As we said before, S_t denotes the underlying fund driving the account value. The dynamics of S_t will be described in the next Section. The account value A_t follows the same dynamics of S_t with the exception of the fact that some fees may be subtracted continuously:

$$dA_t = \frac{A_t}{S_t} dS_t - \alpha_{tot} A_t dt. \tag{2.1}$$

We suppose that total annual fees are charged to the policy holder and withdrawn continuously from the investment account A_t . These fees include the mutual fund management fees α_m and the fee charged to fund the guarantee (also known as the rider) α_g , so that

$$\alpha_{tot} = \alpha_m + \alpha_q.$$

The only portion used by the insurance company to hedge the contract is that coming from α_g : the other fees has to be considered as a outgoing money flow as PH's withdrawals are.

Continuously withdrawn fees are typical of the contract described by Forsyth. Fees may also be withdrawn at the end of each policy year t_i : this is what Kling et al. do in [15]. In this second case

$$dA_t = \frac{A_t}{S_t} dS_t. (2.2)$$

When the PH dies, a death benefit, usually equal to A_t , is payed out to the heirs of the PH. According to the contract's formulation, this death benefit may be payed immediately or at the upcoming event time. If it is payed immediately, the contract stops immediately and the account value and the benefit base becomes equal to zero; otherwise the contract goes on up to the next event time as if nothing happened.

2.4 Event Times

An event time is a sequence of operations under the contract, which occur at fixed dates, usually at each anniversary of the signing of the contract. The times these events take place are denoted by $t_i = \Delta t \cdot i$ and usually $\Delta t = 1$. Let's define $I = T/\Delta t$; then, i runs in $\{0, \ldots I\}$.

When an event time occurs, we assume that the following events happen in this order:

- 1. Withdrawal of the fees by the insurance company (if it is not time continuous);
- 2. If the PH died, payment of the death benefits;

- 3. If the PH is still alive, he (she) is entitled to withdraw a certain amount of money;
- 4. If provided by the contract, a ratchet may increase the benefit base B_t .

We denote with $(A_{t_i}^-, B_{t_i}^-, t_i)$ the state variables just before an event time that occurs at time t_i and with $(A_{t_i}^{k+}, B_{t_i}^{k+}, t_i)$ the state variables just after the update due to the i-th point of the previous numbered list.

2.4.1 Fees

Fees may be withdrawn continuously by the account value, as supposed in Forsyth and Vetzal in[9]. In this case, between two event times, the account value changes as prescribed by (2.1), and nothing special happens at an event time:

$$(A_{t_i}^{1+}, B_{t_i}^{1+}, t_i) = (A_{t_i}^-, B_{t_i}^-, t_i).$$

Otherwise, fees may be withdrawn at the end of the period, as supposed in Kling et al. [15]. In this case, between two event times, the account value changes as prescribed by (2.2), and at the event time, the account value becomes

$$(A_{t_i}^{1+}, B_{t_i}^{1+}, t_i) = (A_{t_i}^- e^{-\alpha_{tot}\Delta t}, B_{t_i}^-, t_i).$$

It is important to be able to deduce the management fees F_t^{man} withdrawn by the account value because they are not used to hedge the contract and therefore they have to be considered as an outgoing money flow. If these fees are withdrawn continuously, we can calculate them observing that their dynamic between two event times is

$$dF_t^{man} = \alpha_m A_t dt + r_t dt.$$

This ODE has the following solution

$$F_t^{man} = \int_0^t e^{\int_s^t r_u du} \alpha_m A_s ds.$$

and can be used in a Monte Carlo approach.

If the fees are withdrawn at the end of the period, we can calculate management fees as a fraction of the total fees withdrawn:

$$F_{t_i}^{tot} = F_{t_{i-1}}^{tot} + A_{t_i}^0 \left(1 - e^{-\alpha_{tot} \Delta t} \right),$$

$$F_{t_i}^{man} = F_{t_{i-1}}^{man} + \frac{\alpha_{man}}{\alpha_{tot}} \left(F_{t_i}^{tot} - F_{t_{i-1}}^{tot} \right).$$

2.4.2 Death Benefit

If the policy holder died at an instant $\bar{t} \in]t_{i-1}, t_i[$ his (her) heirs will obtain a death benefit, that is usually equal to the account value. If the contract provides that the death benefit is paid immediately, then the death benefit DB_t is paid in t and is equal to A_t . Otherwise, if the DB is payed at the next event time, $DB_{t_i} = A_{t_i}^{1+}$ and the contract is concluded (after the DB payment it's worthless):

$$(A_{t_i}^{2+}, B_{t_i}^{2+}, t_i) = (0, 0, t_i).$$

2.4.3 Withdrawal, Bonus, Surrender Event

According to the contract, the policy holder, if still alive at event time t_i , is entitled to withdraw a certain amount W_{t_i} from his (her) police, also if the account value is equal to 0. This amount is given by

$$W_{t_i} = G\Delta t \cdot B_{t_i}^{2+},$$

where G is a constant defined by the contract. In a static framework, we suppose that the PH simply withdraws WA_{t_i} . Otherwise, in a optimization framework, he (she) may withdraw a fraction γ_i of the guaranteed withdrawn:

$$W_{t_i} = \gamma_i G \Delta t \cdot B_{t_i}^{2+}.$$

• The case $\gamma_i = 0$ corresponds to no withdrawal. In this case, the contract may provide a bonus (b_{t_i} is specified by the contract):

$$(A_{t_i}^{3+}, B_{t_i}^{3+}, t_i) = (A_{t_i}^{2+}, B_{t_i}^{2+} (1 + b_{t_i}), t_i).$$

• If $0 < \gamma_i \le 1$ the PH with draws at a lower rate than the standard rate, and the new state variables are

$$(A_{t_i}^{3+}, B_{t_i}^{3+}, t_i) = (\max(0, A_{t_i}^{2+} - W_{t_i}), B_{t_i}^{2+}, t_i).$$

• A third case is possible: the PH may want to withdraw more than the maximum admitted. In this case we suppose $\gamma_i \in]1,2]$, where the case $\gamma_i = 2$ corresponds to a total surrender. We define

$$A' = \max(0, A_{t_i}^{2+} - G\Delta t \cdot B_{t_i}^{2+}).$$

The withdrawn amount is

$$W_{t_i} = G\Delta t \cdot B_{t_i}^{2+} + (\gamma_i - 1) A' (1 - \kappa_{t_i}).$$

where $\kappa_{t_i} \in [0,1]$ is a penalty for withdrawal above the contract amount. The new state variables are

$$(A_{t_i}^{3+}, B_{t_i}^{3+}, t_i) = (\max(0, A_{t_i}^{2+} - G\Delta t \cdot B_{t_i}^{2+} - (\gamma_i - 1) A'), (2 - \gamma_i) B_{t_i}^{2+}, t_i)$$

$$= ((2 - \gamma_i) A', (2 - \gamma_i) B_{t_i}^{2+}, t_i).$$

2.4.4 Ratchet

If the contract species a ratchet (step-up) feature, then the value of the benefit base B is increased if the investment account has increased. The guarantee account B can never decrease, unless the contract is partially or fully surrendered:

$$(A_{t_i}^{4+}, B_{t_i}^{4+}, t_i) = (A_{t_i}^{3+}, \max(B_{t_i}^{3+}, A_{t_i}^{3+}), t_i).$$

Another feature that may be included in the contract is roll-up: for seek of simplicity we won't treat this mechanism.

2.5 Similarity Reduction

An important property of GLWB contract is the fact that these contract behave good under scaling transformations. If $\mathcal{V}(A, B, t)$ denotes the price of a contract, it is possible to prove that for any scalar $\eta > 0$

$$\eta \mathcal{V}(A, B, t) = \mathcal{V}(\eta A, \eta B, t). \tag{2.3}$$

Then, we just have to treat the case $B = \hat{B}$ for a fixed \hat{B} (for example $\hat{B} = P$), and then, choosing $\eta = \hat{B}/B$, we can obtain

$$\mathcal{V}(A,B,t) = \frac{B}{\hat{B}} \mathcal{V}\left(\frac{\hat{B}}{B}A,\hat{B},t\right),\,$$

which means that we can solve the pricing problem only for a single representative value of B. This effectively reduces the problem dimension. The similarity reduction (2.3) was also exploited from Shah et Bertsimas in [19]. We can observe how the reduction similarity works both in the case of a contract that does not contain mechanisms for increasing the base benefits (ratchet), both for contracts with these properties.

3 The Stochastic Models of the Fund S

To understand the different impacts of stochastic volatility and stochastic interest rate over such a long maturity contract, we price the GLWB VA according to two models: the Heston model, which provides stochastic volatility, and the the Black-Scholes Hull-White model, which provide stochastic interest rate. As we said before, the process S represents the underlying fund driving the product's account value A_t .

3.1 The Heston Model

The Heston model [11] is one of the most known and used models in finance to describe the evolution of the volatility of an underlying asset and the underlying asset itself. In order to fix the notation, we report its dynamics:

$$\begin{cases}
dS_t = rS_t dt + \sqrt{V_t} S_t dZ_t^S & S_0 = \bar{S}_0, \\
dV_t = k \left(\theta - V_t\right) dt + \omega \sqrt{V_t} dZ_t^V & V_0 = \bar{V}_0,
\end{cases}$$
(3.1)

where Z^S and Z^V are Brownian motions, and $d\left\langle Z_t^S,Z_t^V\right\rangle = \rho dt.$

3.2 The Black-Scholes Hull-White Model

The Hull-White model [14] is one of historically most important interest rate models, which is nowadays often used for risk-management purposes. The important advantage of the HW model is the existence of the closed formulas for prices of bonds, caplets and swaptions. In order to fix the notation, we report the dynamics of BS HW model:

$$\begin{cases} dS_t = r_t S_t dt + \sigma S_t dZ_t^S & S_0 = \bar{S}_0, \\ dr_t = k \left(\theta_t - r_t \right) dt + \omega dZ_t^r & r_0 = \bar{r}_0, \end{cases}$$

where Z^S and Z^r are Brownian motions, and $d\left\langle Z^S_t, Z^r_t \right\rangle = \rho dt.$

The process r is a generalized Ornstein-Uhlenbeck (hereafter OU) process: here θ_t is not constant but it is a deterministic function which is completely determined by the market values of the zero-coupon bonds by calibration (see Brigo and Mercurio [7]): in this case the theoretical price of ZCB match exactly the market prices.

Let $P^{M}(0,T)$ denote the market price of the zero bond at time 0 for the maturity T. The market instantaneous forward interest rate is then defined by

$$f^{M}\left(0,T\right)=-\frac{\partial\ln P^{M}\left(0,T\right)}{\partial T}.$$

It is well known that the short rate process r can be written as

$$r_t = \omega X_t + \beta(t)$$
.

where X is a stochastic process given by

$$dX_t = -kX_t dt + dZ_t^r, \ X_0 = 0,$$

and $\beta(t)$ is a function

$$\beta\left(t\right)=f^{M}\left(0,t\right)+\frac{\omega^{2}}{2k^{2}}\left(1-\exp\left(-kt\right)\right)^{2}.$$

Then, the BS HW model is described by

$$\begin{cases}
dS_t = r_t S_t dt + \sigma S_t dZ_t^S & S_0 = \bar{S}_0, \\
dX_t = -k X_t dt + dZ_t^T & X_0 = 0, \\
r_t = \omega X_t + \beta(t).
\end{cases}$$
(3.2)

A particular case is called *flat curve*. In this case, we assume $P^{M}\left(t,T\right)=e^{-\bar{r}_{0}\left(T-t\right)}$ and $f^{M}\left(0,T\right)=\bar{r}_{0}$. Then

$$\beta(t) = \bar{r}_0 + \frac{\omega^2}{2k^2} (1 - \exp(-kt))^2,$$

and

$$\theta_t = \bar{r}_0 + \frac{\omega^2}{2k^2} \left(1 - \exp\left(-2kt\right) \right).$$

4 Numerical Methods of Pricing

In this Section we describe the four pricing methods: an Hybrid Monte Carlo method, a standard Monte Carlo method, a Hybrid PDE method, and ADI PDE method.

We remember that our aim is to find the fair value for α_g : it's the value that makes the initial value of the policy equal to the initial gross premium. To achieve this target, we price the police (with one of the following procedures) and then we use the secant method to approach the correct value for α_g . Therefore, the main goal is to be able to find the initial value for a given value of α_g : $V(A_0, B_0, 0)(\alpha_g)$.

We remark that we want to calculate the value of the police from the point of view of the insurance company: the management fees are treated as a outgoing cash flows, and if we assume that the policy holder follows a withdrawal strategy, we consider the worst one for the insurance company.

4.1 The Hybrid Monte Carlo Method

The value of a GLWB police can be calculated through a Monte Carlo set of simulations. This procedure is based in two steps: generation of a scenario (a sampling of the underlying values along the life of the product), and projection of the product into the scenario. According to the way we obtain the scenarios, we distinguish two Monte Carlo models: hybrid MC (HMC) and standard MC (SMC).

The hybrid MC method was introduced in Briani et al. [5]. It is a simple and efficient way to produce MC scenarios for different models. This method is called "hybrid" because it combines trees and MC methods. First, a simple tree needs to be built: this can be done according to Appolloni et al. [2] and also [18], or as we are going to show in 4.1.1. Then, using a vector of Bernoulli random variables, we move from the root through the tree, describing the scenario for the volatility or the interest rate. The values of the underlying at each time step can be easily obtained using an Euler scheme.

4.1.1 Trees

The trees for the Heston model and BS HW model can be obtained from Appolloni et al. [2] or Nelson and Ramaswamy [18]. In this case, the trees are simple binary trees: the node values, and the transition probabilities are set in order to match an approximation of the first two moments of the processe. This kind of tree perform well on short maturity, but the approximation errors accumulate on long maturities. Because of this error that accumulates, the convergence of the algorithm proved to be slow. Therefore, it was necessary to rethink the trees: the main aim was to set up trees which matched exactly some moments of the processes to be diffused. Here we present two trees (see Figure 4.1), one for stochastic volatility and one for stochastic interest rate. They are simple quadrinomial trees, and they are built to match the first 3 moments of the stochastic processes.

We suppose to fix a number N > 0, and we define h = T/N.

The General Case Let Z be a Brownian motion, and let G be a Gaussian stationary process, following

$$dG_t = a(G_t) dt + bdZ_t,$$

with variance that depends only by the time lapse, i.e. $G_{t+s} - G_s | \mathcal{F}_s \sim \mathcal{N}\left(\mu\left(t, G_s\right), \sigma^2\left(t\right)\right)$. We show how to build a simple quadrinomial tree that can match the first three moments.

We define a quadrinomial tree. Let's fix a maturity T, and the number of steps N. Each node will be denoted by $G_{(n,j)}$ where n runs from 0 to N, and j from 0 to 3n. Let h = T/N. The value of each node is

$$G_{(n,j)} = G_0 + (j - 1.5n) \sqrt{\sigma^2(h)}.$$

We remember the first three moments of the process G:

$$M_1 = \mathbb{E}\left[G_{t+h} - G_t | \mathcal{F}_t\right] = \mu\left(h, G_t\right), \ M_2 = \mathbb{E}\left[\left(G_{t+h} - G_t\right)^2 | \mathcal{F}_t\right] = \mu^2\left(h, G_t\right) + \sigma^2\left(h\right),$$

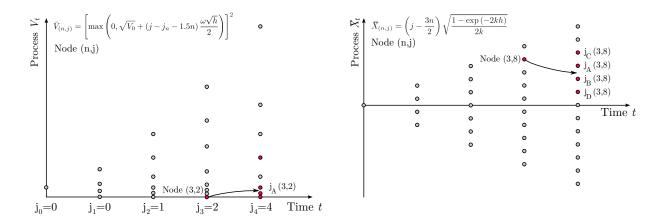


Figure 4.1: The trees for Heston and Hull-White models.

$$M_{3} = \mathbb{E}\left[\left(G_{t+h} - G_{t}\right)^{3} \middle| \mathcal{F}_{t}\right] = \mu^{3}\left(h, G_{t}\right) + 3\mu\left(h, G_{t}\right)\sigma^{2}\left(h\right).$$

Let's fix a node $G_{(n,j)}$. To be brief, μ will denote $\mu(h, G_{(n,j)})$ and σ will denote $\sqrt{\sigma^2(h)}$. We suppose that the expected value μ falls between the values of the nodes at time (n+1)h. This hypothesis can be obtained assuming that the time step h is small enough.

We define

$$j_A\left(n,j\right) = \operatorname{ceil}\left[\frac{G_0 - \mu}{\sigma} + 1.5\left(n+1\right)\right],$$

i.e. the first node in the next time step level whose value is bigger than the mean of the process. This can be seen in Figure 4.1 (both sides): the arrow points out to the expected value of the process, and $j_A(n,j)$ is marked on the Figure. Let

$$j_B(n,j) = j_A(n,j) - 1, \ j_C(n,j) = j_A(n,j) + 1, \ j_B(n,j) = j_D(n,j) - 2.$$

To be brief we will only write j_A, j_B, j_C, j_D , and G_A will be $G_A = G_{(n+1,j_A)}$, and the same for the other letters: this is clear in Figure 4.1, on the right side.

We can now define a Markovian discrete time process \hat{G}_n , n = 0, ..., N with $\hat{G}_0 = G_{(0,0)}$ and we suppose that if $\hat{G}_n = G_{(n,j)}$, then it can move to G_A , G_B , G_C , G_D , according to the following probabilities

$$\begin{split} p_A &= P\left[\hat{G}_{n+1} = G_A | \hat{G}_n = G_{(n,j)}\right] = \frac{(G_A - \mu) \left((G_A - \sigma - \mu)^2 + \sigma^2\right)}{2\sigma^3}, \\ p_B &= P\left[\hat{G}_{n+1} = G_B | \hat{G}_n = G_{(n,j)}\right] = \frac{(\mu - G_A + \sigma) \left((G_A - \mu)^2 + \sigma^2\right)}{2\sigma^3}, \\ p_C &= P\left[\hat{G}_{n+1} = G_C | \hat{G}_n = G_{(n,j)}\right] = \frac{(\mu - G_A + \sigma) \left((G_A - \sigma - \mu)^2 + 2\sigma^2\right)}{6\sigma^3}, \\ p_D &= P\left[\hat{G}_{n+1} = G_D | \hat{G}_n = G_{(n,j)}\right] = \frac{2\sigma^2(G_A - \mu) + (G_A - \mu)^3}{6\sigma^3}. \end{split}$$

And since $G_A - \sigma < \mu \le G_A$, we can easily show that these probabilities are well defined: all in [0,1], their sum is equal to 1, and the first three moments of the variable $\hat{G}_{n+1}|\hat{G}_n = G_{(n,j)}$ are equals to the first three moments of the variable $G_{t+h}|G_t = G_{(n,j)}$.

Now, we approximate the process G by a discrete process \bar{G} that is constant in each time lapse, and is defined as $\bar{G}_t = \hat{G}_{|t/N|}$. The weak convergence of this tree can be proved as in Nelson and Ramaswamy [18].

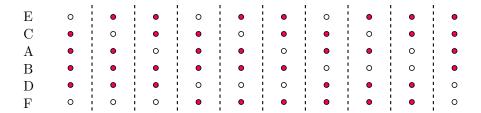


Figure 4.2: The possible combinations used to get positive probabilities in the Heston model tree. The red points correspond to the nodes used.

The Heston Model The Heston process (3.1) for volatility has no constant variance and isn't Gaussian. We consider the process obtained by the square root:

$$d\sqrt{V_t} = \frac{4k\left(\theta - \sqrt{V_t}^2\right) - \omega^2}{8\sqrt{V_t}}dt + \frac{\omega}{2}dZ_t.$$

We approximate it with a Gaussian process with variance $\frac{\omega^2}{4}dt$. This approximation is helpful to define the grid: inspired by [18], we define

$$j_n = \max\left(0, \text{floor}\left(1.5n - \frac{2\sqrt{V_0}}{\omega\sqrt{h}}\right)\right),$$

and we set

$$\bar{V}_{(n,j)} = \left(\max\left(0, \sqrt{V_0} + (j+j_n - 1.5n)\frac{\omega\sqrt{h}}{2}\right)\right)^2.$$

for $j=0,\ldots,3n-j_n$. The shift due to j_n helps to reject the many node with value equal to zero: if $j_n>0$, then $\bar{V}_{(n,0)}=0$ and $\bar{V}_{(n,1)}>0$.

We fix now the value of n and j. The discrete process \bar{V} can jump from a node to another, as in a Markovian chain. We show now how to find the possible upcoming nodes.

The first three moments for the Heston process can be found in Alfonsi [1]:

$$\psi(h) = (1 - e^{-kh})/k, \quad M_1 = \mathbb{E}[V_{t+h}|V_t = v] = ve^{-kh} + \theta k \psi(h),$$

$$M_{2} = \mathbb{E}\left[\left(V_{t+h}\right)^{2} | V_{t} = v\right] = M_{1}^{2} + \omega^{2} \psi\left(h\right) \left[\theta k \psi\left(h\right) / 2 + v e^{-kh}\right],$$

$$M_{3} = \mathbb{E}\left[\left(V_{t+h}\right)^{3} | V_{t} = v\right] = M_{1}M_{2} + \omega^{2}\psi\left(h\right)\left[2v^{2}e^{-2kh} + \psi\left(h\right)\left(k\theta + \frac{\omega^{2}}{2}\right)\left(3ve^{-kh} + \theta k\psi\left(h\right)\right)\right].$$

Then we can proceed as in the general case. Anyway, the grid we're using is based on an approximation: so the probabilities obtained solving the linear system may not be positive.

If we get negative probability for a given node, we try another combination of nodes: the node A or C may be replaced by a node E define as the first node bigger than C, and the node B or D may be replaced with a node F, defined as the smallest before node D. This gives rise to 9 combinations to be tested. If the starting node is small and the node D verifies $j_D = j_n$ we could not do this last change because there would be no F node. In this case we allow the node D to be replaced by the node E: see Figure 4.2.

If these attempts don't give a positive result (negative probabilities), we give up trying to match the first three moments, and we are content to match an approximation of the first two as in [18], thus ensuring the weak convergence. In this case, we only use the nodes A, B, C, D: we define

$$p_{AB} = \frac{\mu - G_{n+1,j_B}}{G_{n+1,j_A} - G_{n+1,j_B}}, \quad p_{BA} = 1 - p_{AB},$$

$$p_{CD} = \frac{\mu - G_{n+1,j_D}}{G_{n+1,j_C} - G_{n+1,j_D}}, \quad p_{DC} = 1 - p_{CD},$$

and

$$p_A = \frac{5}{8} p_{AB}, \ p_B = \frac{5}{8} p_{BA}, \ p_C = \frac{3}{8} p_{CD}, \ p_D = \frac{3}{8} p_{DC}.$$

It is possible to show that the first moment of this variable is equal to M_1 , and as $h \to 0$ the second moment approaches to M_2 , ensuring the convergence, as proved in [18].

In all our numerical tests, this last option (matching only two moments) has never been necessary: changing the nodes, all moments were matched with positive probabilities.

The Hull-White Model The process X in (3.2) is Gaussian. As shown in Ostrovski [20] the variables X_t and $\int_s^t X_y dy$ are bivariate normal distributed conditionally on X_s with well known mean and variance. We define

$$X_{(n,j)} = \left(j - \frac{3n}{2}\right)\sqrt{\frac{1 - \exp\left(-2kh\right)}{2k}}, \ n = 0, \dots, N \text{ and } j = 0, \dots, 3n.$$

Let's fix a node $X_{(n,j)}$. We define

$$H = \exp(-kh), K = \sqrt{\frac{1 - \exp(-2kh)}{2k}}, M_1 = X_{(n,j)}H,$$

$$j_A = \text{ceil}\left[\frac{M_1}{K} + \frac{3(n+1)}{2}\right], \ X_A = X_{(n+1,j_A)}.$$

The transition probabilities are given by

$$\begin{split} p_A &= \frac{(X_A - M_1)}{2K^3} \left(K^2 + \left(K + M_1 - X_A \right)^2 \right), \qquad p_B = \frac{(K + M_1 - X_A)}{2K^3} \left(K^2 + \left(M_1 - X_A \right)^2 \right), \\ p_C &= \frac{(K + M_1 - X_A)}{6K^3} \left(2K^2 + \left(K + M_1 - X_A \right)^2 \right), \qquad p_D = \frac{(X_A - M_1)}{6K^3} \left(2K^2 + \left(M_1 - X_A \right)^2 \right). \end{split}$$

4.1.2 Scenario Generation

The generations of the volatility process and of the interest rate process behave in a similar way: we start from the node (0,0) of the tree and according to a discrete random variable and to the node probabilities, we move to the next node and so on. Let D be a discrete random variable that can assume value A, B, C, D with probabilities p_A, p_B, p_C, p_D : sampling such a variable at each node, we get the values of the process at each time step.

We distinguish two cases for the two models.

The Heston Model We approximate the couple (S_t, V_t) in [0, T] by a discrete process $(\bar{S}_{k\Delta t}, \bar{V}_{k\Delta t})_{k=0,...,T/\Delta t}$ with $(\bar{S}_0, \bar{V}_0) = (S_0, V_0)$. For each scenario, we generate the volatility.

Let $N \sim \mathcal{N}(0,1)$ and $B \sim \mathcal{B}(0.5)$. We deduce the value of $\bar{S}_{t+\Delta t}$ by

$$\bar{S}_{t+\Delta t} = \begin{cases} \bar{S}_{t} \exp\left[\left(r - \frac{\rho}{\sigma}k\theta\right)\Delta t + \left(\frac{\rho}{\sigma}k - \frac{1}{2}\right)\left(\frac{\bar{V}_{t+\Delta t} + \bar{V}_{t}}{2}\right)\Delta t + \frac{\rho}{\sigma}\left(\bar{V}_{t+\Delta t} - \bar{V}_{t}\right) + \sqrt{(1 - \rho^{2})}\Delta t \bar{V}_{t}N\right] & \text{if } B = 0, \\ \bar{S}_{t} \exp\left[\left(r - \frac{\rho}{\sigma}k\theta\right)\Delta t + \left(\frac{\rho}{\sigma}k - \frac{1}{2}\right)\left(\frac{\bar{V}_{t+\Delta t} + \bar{V}_{t}}{2}\right)\Delta t + \frac{\rho}{\sigma}\left(\bar{V}_{t+\Delta t} - \bar{V}_{t}\right) + \sqrt{(1 - \rho^{2})}\Delta t \bar{V}_{t+\Delta t}N\right] & \text{if } B = 1. \end{cases}$$

According to (3.1), we use the normal variable N to generate the Gaussian increment of S, and the Bernoulli variable B to split the operator associated to the Heston process.

This scheme (without splitting) appears in Briani et al. [5] and the splitting method appears in Alfonsi [1].

The Black-Scholes Hull-White Model We approximate the couple (S_t, X_t) in [0, T] by a discrete process $(\bar{S}_{k\Delta t}, \bar{X}_{k\Delta t})_{k=0,...,T/\Delta t}$, with $(\bar{S}_0, \bar{X}_0) = (S_0, 0)$, ad we deduce the interest rate by $\bar{r}_t = \omega \bar{X}_t + \beta(t)$. Let $N \sim \mathcal{N}(0, 1)$. We deduce the value of $\bar{S}_{t+\Delta t}$ by

$$\bar{S}_{t+\Delta t} = \bar{S}_t \exp\left[\left(\frac{\bar{r}_{t\Delta t} + \bar{r}_t}{2} - \frac{\sigma^2}{2}\right) \Delta t + \sigma\left(\left(\bar{X}_{t+\Delta t} + \bar{X}_t \left(k\Delta t - 1\right)\right) \rho + \sqrt{\Delta t} \bar{\rho} N\right)\right].$$

4.1.3 Projection

Once we have generated the scenarios, we project the police into it: it means we calculate the initial value of the contract as the sum of discounted cash flows. This calculation depends on whether we take an optimized strategy or not. Let $\mathcal{V}(A,B,t)$ be the value of a police at time t, having account value equal to A and base benefit equal to B. From now on, we fix a specific scenario. Let V(A,B,t) be the value of a police in that scenario, at time t, having account value equal to A and base benefit equal to B.

Constant Withdrawal In this case the strategy of the PH is fixed: in each event time $\gamma_i = 1$ (for completeness we continue to write γ_i). A simple way to calculate the value of the police is calculating forward the cash flows, conditioning on the death time. As in Holz et al. [12], we have:

$$V\left(A_{0},B_{0},0\right)=\sum_{i=0}^{I}\mathcal{M}\left(t_{i}\right)\left(\sum_{k=0}^{i}e^{-\int_{0}^{t_{k}}r_{s}ds}W_{t_{k}}+e^{-\int_{0}^{t_{i}}r_{s}ds}A_{t_{i}}^{1+}\right).$$

Anyway, we developed another approach, useful for the optimal withdrawal case. First we calculate the values $(A_{t_i}^{4+}, B_{t_i}^{4+}, t_i)$ for all t_i neglecting the effect of mortality (equivalently, assuming that the PH die at the end), with a forward approach:

$$\begin{split} A_{t_i}^{4+} &= \max \left(0, A_{t_{i-1}}^{4+} \frac{S_{t_i}}{S_{t_{i-1}}} e^{-\alpha_{tot}\Delta t} - \gamma_{t_i} G \Delta t B_{t_{i-1}}^{4+} \right), \\ B_{t_i}^{4+} &= \begin{cases} \max \left(B_{t_{i-1}}^{4+}, A_{t_i}^{4+}\right) & \text{if ratchet,} \\ B_{t_{i-1}}^{2+} & \text{otherwise.} \end{cases} \end{split}$$

Then, we proceed backwards, calculating the value of the contract for each time t_i just before the withdrawal. The value of the contract at time t_i can be written as the discounted value at time t_{i+1} plus the discounted value of the cash flows relating the period $\begin{bmatrix} t_i^{4+}, t_{i+1}^{4+} \end{bmatrix}$. The final condition on the value of the contract is

$$V\left(A_T^{4+},B_T^{4+},T\right)=0,$$

because all PHs are death and all benefits have been paid. Then

$$V\left(A_{t_{i}}^{4+},B_{t_{i}}^{4+},t_{i}\right)=e^{-\int_{t_{i}}^{t_{i+1}}r_{s}ds}\left[V\left(A_{t_{i+1}}^{4+},B_{t_{i+1}}^{4+},t_{i+1}\right)+\mathcal{R}\left(t_{i+1}\right)W_{t_{i+1}}\right]+DB+MF,$$

where DB and MF stand for the discounted value in t_i of the death benefit and management fees paid in $\begin{bmatrix} t_i^{4+}, t_{i+1}^{4+} \end{bmatrix}$. We distinguish four cases depending on how the management fees and the death benefit are payed.

CASE 1: DB payed at the end, Fees withdrawn at the end

$$DB = \mathcal{M}\left(t_{i}\right)e^{-\int_{t_{i}}^{t_{i+1}}r_{s}ds}A_{t_{i}}^{4+\frac{S_{t_{i+1}}}{S_{t_{i}}}}e^{-\alpha_{tot}\Delta t}, \qquad MF = \mathcal{R}\left(t_{i}\right)e^{-\int_{t_{i}}^{t_{i+1}}r_{s}ds}A_{t_{i}}^{4+\frac{S_{t_{i+1}}}{S_{t_{i}}}}\left(1-e^{-\alpha_{tot}\Delta t}\right)\frac{\alpha_{m}}{\alpha_{tot}}$$

CASE 2: DB payed at the end, Fees withdrawn continuously

$$DB = \mathcal{M}\left(t_{i}\right) e^{-\int_{t_{i}}^{t_{i+1}} r_{s} ds} A_{t_{i}}^{4+\frac{S_{t_{i+1}}}{S_{t_{i}}}} e^{-\alpha_{tot} \Delta t}, \qquad MF = \mathcal{R}\left(t_{i}\right) \alpha_{m} \frac{A_{t_{i}}^{4+}}{S_{t_{i}}} \int_{t_{i}}^{t_{i+1}} e^{-\int_{t_{i}}^{t} r_{s} ds} S_{t} e^{-\alpha_{tot}(t-t_{i})} dt.$$

CASE 3: DB payed immediately, Fees withdrawn at the end

$$DB = \mathcal{M}(t_i) \frac{A_{t_i}^{4+}}{S_{t_i}} \int_{t_i}^{t_{i+1}} e^{-\int_{t_i}^{t} r_s ds} S_t e^{-\alpha_{tot}(t-t_i)} dt,$$

$$MF = \mathcal{M}\left(t_{i}\right) \frac{\alpha_{m}}{\alpha_{tot}} \frac{A_{t_{i}}^{4+}}{S_{t_{i}}} \int_{t_{i}}^{t_{i+1}} S_{t}\left(1 - e^{-\alpha_{tot}(t - t_{i})}\right) e^{-\int_{t_{i}}^{t} r_{u} du} dt + \mathcal{R}\left(t_{i+1}\right) e^{-\int_{t_{i}}^{t_{i+1}} r_{s} ds} A_{t_{i}}^{4+} \frac{S_{t_{i+1}}}{S_{t_{i}}} \left(1 - e^{-\alpha_{tot}\Delta t}\right) \frac{\alpha_{m}}{\alpha_{tot}}.$$

CASE 4: DB payed immediately, Fees withdrawn continuously

$$DB = \mathcal{M}(t_i) \frac{A_{t_i}^{4+}}{S_{t_i}} \int_{t_i}^{t_{i+1}} e^{-\int_{t_i}^{t} r_s ds} S_t e^{-\alpha_{tot}(t-t_i)} dt,$$

$$MF = \mathcal{M}(t_i) \alpha_m \frac{A_{t_i}^{4+}}{S_{t_i}} \int_{t_i}^{t_{i+1}} S_t e^{-\alpha_{tot}(t-t_i)} e^{-\int_{t_i}^t r_u du} (t_{i+1} - t) dt + \mathcal{R}(t_{i+1}) \alpha_m \frac{A_{t_i}^{4+}}{S_{t_i}} \int_{t_i}^{t_{i+1}} e^{-\int_{t_i}^t r_s ds} S_t e^{-\alpha_{tot}(t-t_i)} dt.$$

Proceeding in this way, it is possible to calculate $V(A_0^{4+}, B_0^{4+}, 0)$. The initial value of the police is

$$V(A_0^-, B_0^-, 0) = V(A_0^{4+}, B_0^{4+}, 0)$$

if the first withdrawal takes place at time $t = t_1$, or

$$V(A_0^-, B_0^-, 0) = V(A_0^{4+}, B_0^{4+}, 0) + \gamma_0 G \Delta t P$$

if the first withdrawal takes place at time t=0. Then we simply have to calculate the average of $V\left(A_0^-, B_0^-, 0\right)$ among the simulated scenarios.

Optimal Withdrawal In this case we suppose that at each event time t_i the policy holder can withdraw a fraction γ_i of the regular amount. To price in this case, we suppose that the PH chooses the value of γ that causes the worst hedging case for the insurance company. We denote $\mathcal{V}(A, B, t)$ the expected value at time t of a generic police whose state parameters are A, B:

$$V(A, B, t) = \mathbb{E}[V(A, B, t)].$$

So, we suppose that the policy holder chooses γ_i such that

$$\gamma_i = \underset{\gamma \in [0,2]}{\operatorname{argmax}} \mathcal{V}\left(A^{4+}, B^{4+}, t\right).$$

This expected value can be calculated with a Longstaff-Schwartz approach:

- 1. Simulate N random scenarios and price the police into these scenarios using random values for γ_i .
- 2. For t = T to t = 0:
 - (a) Approximate the function $\mathcal{V}(A, B, t)$ using the least squares projection into a space of functions (usually polynomials).
 - (b) For each scenario find the optimal value of γ_t .
 - (c) Recalculate the upcoming state variables from s = t to s = T assuming that the policy holder chooses the best value for γ .
- 3. Calculate the average of the initial value $V(A_0, B_0, 0)$ for all the scenarios.

The approximation of the function $\mathcal{V}(A,B,t)$ can be improved by the reduction property.

4.2 Standard Monte Carlo Method

The Monte Carlo method is very similar to the hybrid Monte Carlo one. The only different thing, is the way we produce the random scenarios. The projection phase is the same as hybrid Monte Carlo's one.

4.2.1 Scenario generation

We distinguish two cases for the two models.

The Heston Model The generation of the scenarios (underlying and volatility) in this case has been done using a third order schemes described in Alfonsi [1].

The Black-Scholes Hull-White Model The generation of the scenarios (underlying and interest rate) in this case has been done using an exact schemes described in Ostrovski [20], with a few changes in order to incorporate the correlation between underlying and interest rate.

4.3 PDE Hybrid Method

The Hybrid PDE approach is different from the previous ones. In fact it's a PDE pricing method and it's based on Briani et al. [4],[5] both for Heston and Hull-White case. Using a tree to diffuse volatility or interest rate, we freeze these values between two tree-levels and we solve four PDE for each tree node; then we mix the values given by the four PDEs according to the transition probabilities of each node.

We can resume the pricing methods in three features: model, algorithm structure and pricing.

4.3.1 The Heston Model

Starting from the model

$$\begin{cases} dS_t = rS_t dt + \sqrt{V_t} S_t \left(\rho dZ_t^V + \bar{\rho} dZ_t^S \right) & V_0 = \bar{V}_0, \\ dV_t = k \left(\theta - V_t \right) dt + \omega \sqrt{V_t} dZ_t^V & S_0 = \bar{S}_0, \end{cases} \quad d \left\langle Z_t^S, Z_t^V \right\rangle = 0,$$

we define the process

$$E_{t} = \ln(A_{t}) - \frac{\rho}{\omega} V_{t}, \ E_{0} = \ln(A_{0}) - \frac{\rho}{\omega} V_{0},$$

$$A_{t} = \exp\left(E_{t} + \frac{\rho}{\omega} V_{t}\right). \tag{4.1}$$

Then

$$dE_{t} = \left(r - \frac{V_{t}}{2} - \frac{\rho}{\omega}k\left(\theta - V_{t}\right) - \alpha_{tot}\right)dt + \sqrt{\left(1 - \rho^{2}\right)V_{t}}dZ_{t}^{S},$$

if fees are taken continuously, otherwise

$$dE_t = \left(r - \frac{V_t}{2} - \frac{\rho}{\omega}k\left(\theta - V_t\right)\right)dt + \bar{\rho}\sqrt{\left(1 - \rho^2\right)V_t}dZ_t^S.$$

4.3.2 The Black-Scholes Hull-White Model

Starting from the model

$$\begin{cases} dS_t = r_t S_t dt + \sigma S_t \left(\rho dZ_t^r + \bar{\rho} dZ_t^S \right) & S_0 = \bar{S}_0, \\ dX_t = -k X_t dt + dZ_t^r & X_0 = 0, \quad d \left\langle Z_t^S, Z_t^r \right\rangle = 0, \\ r_t = \omega X_t + \beta \left(t \right), \end{cases}$$

we define the process

$$U_t = \ln(A_t) - \rho \sigma X_t, \ U_0 = \ln(A_0),$$

$$A_t = \exp(U_t + \rho \sigma X_t).$$
(4.2)

Then

$$dU_t = \left(r_t - \frac{\sigma^2}{2} + \sigma \rho k X_t - \alpha_{tot}\right) dt + \sigma \sqrt{1 - \rho^2} dZ_t^S,$$

if fees are taken continuously, otherwise

$$dU_t = \left(r_t - \frac{\sigma^2}{2} + \sigma \rho k X_t\right) dt + \sigma \sqrt{1 - \rho^2} dZ_t^S.$$

4.3.3 Algorithm structure

The structures for this algorithm consist in a tree and a PDE solver. As described in Briani et al. [4],[5], we use a tree to diffuse the volatility (or the interest rate) along the life of the product, and we solve backward a 1D PDE freezing at each node of the tree the volatility (or the interest rate). The tree is built according to Section 4.1.1 (quadrinomial tree, matching the first three moments of the process), and the PDE is solved with a finite difference approach. We have to solve the PDE between event time, and at the event time we apply the changes to the states to reproduce the effects of the events.

4.3.4 Pricing

The PDE we have to solve is essentially the same as in Forsyth and Vetzal [9]. We distinguish four cases as we did in Monte Carlo case. We denote with $\mathcal{V}(A,B,t)$ the value of a contract at time t, whose account value is worth A and whose base benefit is worth B. Consequently, we define

$$\mathcal{V}^{He}\left(E,B,t\right) = \mathcal{V}\left(\exp\left(E + \frac{\rho}{\omega}V_{t}\right),B,t\right),$$

and

$$\mathcal{V}^{HW}(U, B, t) = \mathcal{V}(\exp(U + \rho \sigma X_t) B, t).$$

The variables \bar{r} , \bar{X} and \bar{V} will denote the frozen values of r_t , X_t and V_t . We solve the transformed PDE between two event times for each node of the tree four times: one for each of the possible next nodes, using the initial data corresponding to these nodes. To reduce the run time, we do this only for most relevant nodes: this cutting technique dramatically reduced calculation times without compromising the quality of results. Then, using the inverse transformations (4.1) and (4.2), we apply the event times' actions. In the next few paragraphs, we are going to write 2 PDEs: one for the Heston model, and one for the BS HW model.

CASE 1: DB payed at the end, Fees withdrawn at the end

The terminal condition is

$$\mathcal{V}(A, B, T) = \mathcal{R}(T - \Delta t) A \left(1 - \left(1 - e^{-\alpha_{tot} \Delta t} \right) \frac{\alpha_g}{\alpha_{tot}} \right).$$

The associated PDEs are

$$\mathcal{V}_{t}^{He} + \frac{\bar{\rho}^{2}\bar{V}}{2}\mathcal{V}_{EE}^{He} + \left(r - \frac{\bar{V}}{2} - \frac{\rho}{\omega}k\left(\theta - \bar{V}\right)\right)\mathcal{V}_{E}^{He} - r\mathcal{V}^{He} = 0, \tag{He 1}$$

$$\mathcal{V}_t^{HW} + \frac{\bar{\rho}^2 \sigma^2}{2} \mathcal{V}_{UU}^{HW} + \left(\bar{r} - \frac{\sigma^2}{2} + \sigma \rho k \bar{X}\right) \mathcal{V}_U^{HW} - \bar{r} \mathcal{V}^{HW} = 0.$$
 (HW 1)

For $t_i = T - 1$ to $t_i = 0$ we have to:

- 1. Solve the PDE backward from t_{i+1} to t_i .
- 2. Calculate the value of \mathcal{V} from the value of \mathcal{V}^{He} or \mathcal{V}^{HW} .
- 3. In case of ratchet $\mathcal{V}(A, B, t_i^{3+}) = \mathcal{V}(A, \max(A, B), t_i^{4+})$
- 4. Withdrawal:

(a) if
$$\gamma_{t_i} = 0$$
: $\mathcal{V}(A, B, t_i^{2+}) = \mathcal{V}(A, B(1 + b_{t_i}), t_i^{3+})$;

(b) if $\gamma_{t_i} \in [0,1]$:

$$\mathcal{V}\left(A, B, t_i^{2+}\right) = \mathcal{V}\left(\max\left(0, A - \gamma_{t_i} G \Delta t B\right), B, t_i^{3+}\right) + \mathcal{R}\left(t_i\right) \gamma_{t_i} G \Delta t B;$$

(c) if $\gamma_{t_i} \in]1, 2]$:

$$\mathcal{V}\left(A,B,t_{i}^{2+}\right) = \mathcal{V}\left(\max\left(0,A-G\Delta tB\right)\left(2-\gamma_{t_{i}}\right),B\left(2-\gamma_{t_{i}}\right),t_{i}^{3+}\right) + \\ + \mathcal{R}\left(t_{i}\right)\left(G\Delta tB + \left(\gamma_{t_{i}}-1\right)\max\left(0,A-G\Delta tB\right)\left(1-\kappa_{t_{i}}\right)\right);$$

- 5. Death benefit: $\mathcal{V}\left(A,B,t_{i}^{1+}\right) = \mathcal{V}\left(A,B,t_{i}^{2+}\right) + \left(\mathcal{R}\left(t_{i-1}\right) \mathcal{R}\left(t_{i}\right)\right)A$.
- 6. Fees: $V\left(A, B, t_i^-\right) = V\left(Ae^{-\alpha_{tot}\Delta t}, B, t_i^{1+}\right) + \mathcal{R}\left(t_{i-1}\right) \frac{\alpha_m}{\alpha_{tot}} A\left(1 e^{-\alpha_{tot}\Delta t}\right)$.
- 7. Calculate the value of \mathcal{V}^{He} or \mathcal{V}^{HW} from the value of \mathcal{V} .

CASE 2: DB payed at the end, Fees withdrawn continuously

The differences between this case and the case 1 are the following ones. The terminal condition is

$$V(A, B, T) = \mathcal{R}(T - \Delta t) A.$$

The associated PDEs are

$$\mathcal{V}_{t}^{He} + \frac{\bar{\rho}^{2}\bar{V}}{2}\mathcal{V}_{EE}^{He} + \left(r - \frac{\bar{V}}{2} - \frac{\rho}{\omega}k\left(\theta - \bar{V}\right) - \alpha_{tot}\right)\mathcal{V}_{E}^{He} - r\mathcal{V}^{He} + \alpha_{m}\mathcal{R}\left(t\right)\exp\left(E_{t} + \frac{\rho}{\omega}\bar{V}\right) = 0, \quad (\text{He 2})$$

$$\mathcal{V}_{t}^{HW} + \frac{\bar{\rho}^{2}\sigma^{2}}{2}\mathcal{V}_{UU}^{HW} + \left(\bar{r} - \frac{\sigma^{2}}{2} + \sigma\rho k\bar{X} - \alpha_{tot}\right)\mathcal{V}_{U}^{HW} - \bar{r}\mathcal{V}^{HW} + \alpha_{m}\mathcal{R}\left(t\right)\exp\left(U_{t} + \rho\sigma\bar{X}\right) = 0. \tag{HW 2}$$

Point 6 (fees step) becomes

$$\mathcal{V}\left(A,B,t_{i}^{-}\right)=\mathcal{V}\left(A,B,t_{i}^{1+}\right).$$

CASE 3: DB payed immediately, Fees withdrawn at the end

The differences between this case and the case 1 are the following ones. The terminal condition is

$$\mathcal{V}(A, B, T) = 0.$$

The associated PDEs are

$$\mathcal{V}_{t}^{He} + \frac{\bar{\rho}^{2}\bar{V}}{2}\mathcal{V}_{EE}^{He} + \left(r - \frac{\bar{V}}{2} - \frac{\rho}{\omega}k\left(\theta - \bar{V}\right)\right)\mathcal{V}_{E}^{He} - r\mathcal{V}^{He} + \mathcal{M}\left(t_{i}\right)\exp\left(E_{t} + \frac{\rho}{\omega}\bar{V}\right)\left(1 - \left(1 - e^{-\alpha_{tot}(t - t_{i})}\right)\frac{\alpha_{g}}{\alpha_{tot}}\right) = 0 , \text{ (He 3)}$$

$$\mathcal{V}_{t}^{HW} + \frac{\bar{\rho}^{2}\sigma^{2}}{2}\mathcal{V}_{UU}^{HW} + \left(\bar{r} - \frac{\sigma^{2}}{2} + \sigma\rho k\bar{X}\right)\mathcal{V}_{U}^{HW} - \bar{r}\mathcal{V}^{HW} + \mathcal{M}\left(t_{i}\right)\exp\left(U_{t} + \rho\sigma\bar{X}\right)\left(1 - \left(1 - e^{-\alpha_{tot}(t - t_{i})}\right)\frac{\alpha_{g}}{\alpha_{tot}}\right) = 0. \text{ (HW 3)}$$

Point 5 (death benefit step) and 6 (fees step) become:

- Death benefit: $\mathcal{V}(A, B, t_i^{1+}) = \mathcal{V}(A, B, t_i^{2+})$.
- Fees: $\mathcal{V}\left(A, B, t_i^-\right) = \mathcal{V}\left(Ae^{-\alpha_{tot}\Delta t}, B, t_i^{1+}\right) + \mathcal{R}\left(t_i\right) \frac{\alpha_m}{\alpha_{tot}} A\left(1 e^{-\alpha_{tot}\Delta t}\right)$.

CASE 4: DB payed immediately, Fees withdrawn continuously

The differences between this case and the case 1 are the following ones. The terminal condition is

$$\mathcal{V}\left(A,B,T\right)=0.$$

The associated PDEs are

$$\mathcal{V}_{t}^{He} + \frac{\bar{\rho}^{2}\bar{V}}{2}\mathcal{V}_{EE}^{He} + \left(r - \frac{\bar{V}}{2} - \frac{\rho}{\omega}k\left(\theta - \bar{V}\right) - \alpha_{tot}\right)\mathcal{V}_{E}^{He} - r\mathcal{V}^{He} + \exp\left(E_{t} + \frac{\rho}{\omega}\bar{V}\right)\left(\alpha_{m}\mathcal{R}\left(t\right) + \mathcal{M}\left(t_{i}\right)\right) = 0 , \quad (\text{He 4})$$

$$\mathcal{V}_{t}^{HW} + \frac{\bar{\rho}^{2}\sigma^{2}}{2}\mathcal{V}_{UU}^{HW} + \left(\bar{r} - \frac{\sigma^{2}}{2} + \sigma\rho k\bar{X} - \alpha_{tot}\right)\mathcal{V}_{U}^{HW} - \bar{r}\mathcal{V}^{HW} + \exp\left(U_{t} + \rho\sigma\bar{X}\right)\left(\alpha_{m}\mathcal{R}\left(t\right) + \mathcal{M}\left(t_{i}\right)\right) = 0. \text{ (HW 4)}$$

Point 5 (death benefit step) and 6 (fees step) become

$$\mathcal{V}\left(A, B, t_i^-\right) = \mathcal{V}\left(A, B, t_i^{1+}\right) = \mathcal{V}\left(A, B, t_i^{2+}\right).$$

This concludes the static withdrawal case. In the optimal withdrawal case, we suppose the PH to change the value of γ_i used in step n. 4 (withdrawal step). He (she) will choose the value of $\gamma_i \in [0,2]$ in order to maximizes the value of $\mathcal{V}\left(A,B,t_i^{2+}\right)$. This maximization can be done using a grid of values for γ_i and choosing at each time the best value.

4.4 PDE ADI Method

Consider the asset price process given by the system of stochastic differential equations described in Section 2. We describe the ADI method only in the case 2, but the other cases can be easily adapted. Moreover, we have chosen to not use the transformed PDE described in Section 4.3.4, but the classical version of PDEs for the Black-Scholes, Heston and Black-Scholes Hull-White model. The associated PDEs are

$$\mathcal{V}_{t}^{He} + \frac{VA^{2}}{2}\mathcal{V}_{AA}^{He} + \frac{\omega^{2}V}{2}\mathcal{V}_{VV}^{He} + (r - \alpha_{tot})A\mathcal{V}_{A}^{He} + \rho\omega AV\mathcal{V}_{AV}^{He} + k(\theta - V)\mathcal{V}_{V}^{He} - r\mathcal{V}^{He} + \alpha_{m}\mathcal{R}(t)A = 0 \quad \text{(He 2b)}$$

$$\mathcal{V}_{t}^{HW} + \frac{\sigma^{2}A^{2}}{2}\mathcal{V}_{AA}^{HW} + \frac{\omega^{2}}{2}\mathcal{V}_{rr}^{HW} + (r - \alpha_{tot})A\mathcal{V}_{A}^{HW} + \rho\omega A\sigma\mathcal{V}_{Ar}^{HW} + k\left(\theta_{t} - r\right)\mathcal{V}_{r}^{HW} - r\mathcal{V}^{HW} + \alpha_{m}\mathcal{R}\left(t\right)A = 0 \quad (HW 2b)$$

Because of the long maturity, solving a two-dimensional PDE is a very costly and slow method. The idea is to use splitting schemes of ADI (alternating directional implicit) type. In this paper, we only present the Douglas scheme, but various scheme are available in the literature. In order to solve the PDE, we should address many numerical difficulties. The first one is the mesh and we have chosen to use the meshes described in [13] with the parameters

$$A_{left} = 0.8S_0$$
 $A_{right} = 1.2S_0$ $A_{max} = 100S_0$ and $d_1 = S_0/20$,

for the mesh of variable A,

$$R_{\text{max}} = 10R_0$$
, $c = R_0$ and $d_2 = R_{\text{max}}/400$

for the mesh of variable r in the Black-Scholes Hull-White model, and

$$V_{\text{max}} = MIN(MAX(100V_0, 1), 5)$$
 and $d_3 = V_{\text{max}}/500$.

for the mesh of variable V in the Heston model. The second difficulty is the choice of the splitting scheme. We have chosen the Douglas scheme with parameter $\theta = 1/2$ because it is the easiest to implement, but of

course some higher order schemes (in time) would be more optimal. The last difficulty, but not the least, is the choice of boundary conditions. Since there is no closed form solutions for the GLWB product, it is difficult to make the right choice for the boundary conditions. Moreover the boundary conditions have a big impact on the solution, because of the long maturity. The choice of homogeneous Neumann conditions is usually done because it simplifies the system to solve (exactly it simplifies the finite difference scheme at the boundary). In the context of GLWB, the boundary conditions for the Black-Scholes Hull-White model will be given by:

$$\begin{split} \frac{\partial \mathcal{V}_t^{HW}}{\partial s}(A,r,t) &= 0, \qquad \text{if } A = 0 \text{ or } A = A_{\max}, \\ \frac{\partial \mathcal{V}_t^{HW}}{\partial v}(A,r,t) &= 0, \qquad \text{if } r = \pm R_{\max}, \end{split}$$

on the mesh $[0, A_{\text{max}}] \times [-R_{\text{max}}, R_{\text{max}}]$, and the boundary condition for the Heston model will be given by:

$$\begin{split} &\frac{\partial \mathcal{V}_{t}^{He}}{\partial s}(A,V,t)=0, \qquad \text{if } A=0 \text{ or } A=A_{\max}, \\ &\frac{\partial \mathcal{V}_{t}^{He}}{\partial v}(A,V,t)=0, \qquad \text{if } V=V_{\max}, \end{split}$$

on the mesh $[0, A_{\text{max}}] \times [0, V_{\text{max}}]$, and with no condition at V = 0 since it is an outflow boundary.

5 Numerical Results

In this Section we compare the numerical methods used in Section 4: Hybrid Monte Carlo (HMC), Standard Monte Carlo (SMC), Hybrid PDE (HPDE), and ADI PDE (APDE). In particular we compare pricing and Greeks computation in $Static\ Case\ 5.1$ and $Dynamic\ Case\ 5.2$.

We chose the parameters of the methods according to 4 configurations (A, B, C, D), with an increasing number of steps, and so that the calculation time for the various methods in each configuration were close. There 4 configurations are in Table 1, with the notation (time steps per year; space steps; vol steps) for the ADI PDE method, (time steps per year; space steps) for the Hybrid PDE method approaches, and (time steps per year; number of simulations) for the MC's one. In Monte Carlo for dynamic case, we also add the degree of the approximating polynomial. These values had been chosen to achieve approximately these run times: (A) 30 s, (B) 120 s, (C) 480 s, (D) 1900 s. To reduce the run time we do the secant iterations using an increasing number of time steps for all the methods: the values in Table 1 are those used for the last 3 iterations.

We use the standard MC both as a pricing method, both as a benchmark (BM). About the benchmark, in the static case we used 10^7 independent runs. In the dynamic case we used 10^6 independent runs, arranged in 10 sub runs; in each sub runs the expected value has been approximated by a 6 order polynomial. At each event time, the PH can chose between $\gamma = 0$, $\gamma = 1$ and $\gamma = 2$.

The search for the fair α_g value has been driven by the secant method. The initial values for this method were $\alpha_g = 0$ bp and $\alpha_g = 200$ bp.

To achieve Delta calculation in Monte Carlo methods we used a 1% shock in static case and 1% in dynamic case.

We used the DAV 2004R mortality Table, 65 year old German male (see [9] for the Table). It contains the probabilities that a person aged t will die within the next year. It's easy to get the function \mathcal{M} from these probabilities.

5.1 Static Case

In the static case we suppose the PH to withdrawal exactly at the guaranteed rate: $\gamma_t = 1$.

		BS HW	Static		HESTON STATIC				
	нмс	$_{\mathrm{SMC}}$	$_{ m HPDE}$	APDE	HMC	$_{\mathrm{SMC}}$	$_{ m HPDE}$	APDE	
A	$5 \times 1.3 \cdot 10^5$	$1\!\times\!2.7\!\cdot\!10^5$	30×400	$18 \times 180 \times 36$	$5 \times 8.6 \cdot 10^4$	$5\!\times\!7.4\!\cdot\!10^4$	35×400	$26 \times 260 \times 13$	
B	$3 \mid 10 \times 2.3 \cdot 10^5 1 \times 9.8 \cdot 10^5 60 \times 600$		$27\!\times\!270\!\times\!54$	$10\!\times\! 1.6\!\cdot\! 10^{5}$	$10\!\times\!1.4\!\cdot\!10^5$	70×600	$40\!\times\!400\!\times\!20$		
C	$20\!\times\!5.4\!\cdot\!10^5$	$1\!\times\!4.9\!\cdot\!10^6$	$100\!\times\!1000$	$40\!\times\!400\!\times\!80$	$20\!\times\!3.8\!\cdot\!10^5$	$20\!\times\!3.5\!\cdot\!10^5$	$100\!\times\!1000$	$64\!\times\!640\!\times\!32$	
D	$40 \times 1.0 \cdot 10^6$	$1\!\times\!2.0\!\cdot\!10^7$	$200\!\times\!2000$	$62 \times 620 \times 124$	$40\!\times\!7.3\!\cdot\!10^5$	$40\!\times\!7.5\!\cdot\!10^{5}$	$200\!\times\!2000$	$104 \times 1040 \times 52$	
		BS HW 1	DYNAMIC			HESTON	DYNAMIC		

		BS HW I	DYNAMIC		HESTON DYNAMIC					
	HMC SMC HPDE APDE		HMC	$_{\mathrm{SMC}}$	$_{ m HPDE}$	APDE				
A	$5 \times 3.3 \cdot 10^3 \times 2$	$5\times3.2\cdot10^3\times2$	30×400	$16\!\times\!160\!\times\!32$	$5 \times 3.2 \cdot 10^3 \times 2$	$5 \times 3.2 \cdot 10^3 \times 2$	35×400	$22\!\times\!220\!\times\!11$		
B	$10 \times 1.610^4 \times 3$	$5 \times 1.6 \cdot 10^4 \times 3$	60×600	$24\!\times\!240\!\times\!48$	$10 \times 1.510^4 \times 3$	$10 \times 1.510^4 \times 3$	70×600	$36\!\times\!360\!\times\!18$		
C	$20 \times 5.2 \cdot 10^4 \times 4$	$5 \times 5.3 \cdot 10^4 \times 4$	$100\!\times\!1000$	$38\!\times\!380\!\times\!76$	$20 \times 4.910^4 \times 4$	$20 \times 4.910^4 \times 4$	$100\!\times\!1000$	$60\!\times\!600\!\times\!30$		
D	$40 \times 1.410^5 \times 5$	$5 \times 1.6 \cdot 10^5 \times 5$	$200\!\times\!2000$	$60{\times}600{\times}120$	$40 \times 1.310^5 \times 5$	$40 \times 1.310^5 \times 5$	$200\!\times\!2000$	$100 \times 1000 \times 50$		

Table 1: Configuration parameters for the BS HW model and for the Heston model, static and dynamic.

The static tests 1 and 2 are inspired by [9]: in their article, Forsyth and Vetzal price a GLWB contract in a static framework, under the Black Scholes model with r = 0.04 and $\sigma = 0.15$. The contract parameters are reported in the Table 2; the contract type corresponds to case 2 in Section 4.1.3 and 4.3.4.

Initial age of PH	65	Gr. premium	100	DB payment	next anniv.
G	0.05	Initial fees	0	Ratchet	Off/On (annual)
Withdrawal rate	1 per Y	α_m	0	Strategy	static $(\gamma = 1)$
First withdrawal	1^{st} anniv.	Fees taken	$\operatorname{cont.ly}$		

Table 2: The contract parameters for static tests (except Test 2B).

They treated two cases: no ratchet, and annual ratchet. In the first case they get $\alpha_g = 35.51$ bp and in the second case $\alpha_g = 64.92$ bp. In Test 1 and Test 2 we introduce respectively stochastic interest rate and stochastic volatility to analyze the impact of these model developments on the fair guarantee fee. The parameters for interest rate and volatility models has been chosen to be plausible.

To compare our results in the Heston model with Kling's ones in [15] we performed test 2B. In this case, product parameters are reported in Table 3, and correspond to case 1 in Section 4.1.3 and 4.3.4.

Initial age of PH	65	Gr. premium	100	DB payment	next anniv.
G	$4.90\%, \ 4.19\% \ \text{if ratchet}$	Initial fees	4%	Ratchet	Off/On (annual)
Withdrawal rate	1 per Y	α_m	151 bp	Strategy	static $(\gamma = 1)$
First withdrawal	1^{st} anniv.	Fees taken	at the end		

Table 3: The contract parameters for Test 2B-Static.

5.1.1 Test 1-Static: the Black-Scholes Hull-White Model

In this test we want to price a product according to BS HW model. We use the same corresponding parameters as in test [9]. Model parameters are shown in the Table 4. Results are available in Table 5.

All four methods behave well and in the configuration D, gave results consistent with the benchmark. HPDE proved to be the best: all configurations gave results consistent with the benchmark. Then APDE and SMC, and HMC gave good results too. SMC performed a little better than HMC: the first method simulates the underlying value and the interest rate exactly and so it is enough to simulate the values at each event time. HMC matches the first three moments of the BS HW r process, but doesn't reproduce exactly its law: therefore it is right to increase the number of steps per year. So, for a given run time, we can simulate less scenarios in HMC than SMC: effectively, the confidence interval of HMC is larger than SMC's

one. Moreover, SMC over performed the benchmark when using configuration D. Particularly, correlation between underlying and interest rate has a fundamental role, and it's impact can be bigger than ratchet's impact: for example, case no ratchet with $\rho = 0.5$ gave a higher price than case ratchet with $\rho = -0.5$ (111 bp vs 84 bp).

5.1.2 Test 2-Static: the Heston Model

In this test we want to price a product according to the Heston model. Model parameters are shown in the Table 6. Results are shown in Table 7.

In this Test, MC methods had more problems; PDE methods' values are close to the benchmark, while MC method's values were far, but compatibles with the benchmark (BM's value is inside MCs' confidence interval). Probably, in this case, the benchmark is not very accurate: this is due to the fact we used SMC to calculate it. If we compare the two MC approaches, in this case, they both use a third order approximation and than they become equivalent: HMC proved to be faster than SMC when using few time steps (we could exploit +16% simulations in configuration A), while SMC proved to be slightly faster in high time steps simulations, because of more time needed to build the volatility tree (-3% simulations in configuration D). HPDE showed to be very stable (case no ratchet, $\rho = -0.5$, α_g didn't changed through configurations B-D), but APDE behaved well to (monotone convergence). In the Heston model, correlation has a less important role than in BS HW case: among the different values of ρ , the value of α_g changes less then 5 bp in no-ratchet case, and less than 1.5 bp in ratchet case.

5.1.3 Test 2B-Static: the Heston Model

In this test we want to obtain the results shown in [15], where the contract are priced with MC techniques. The values given in [15] are 150 bp for both cases (no ratchet and ratchet). Model parameters are given in Table 8. Results are available in Table 9.

In this Test, all methods gave the same results, but not the same results as in Kling et al. [15]. One possibility is that we have misinterpreted some of the contractual specifications in Kling's paper, leading to some subtle differences in the contracts that we are considering as compared to theirs, and these discrepancies result in different fees. Another potential explanation is that a Monte Carlo method was used to determine the fee by Kling et al.; this may have introduced a significant error when calculating the fee unless a very large number of simulations was used. They didn't report a confidence interval for their results, so it's hard to understand the cause of the gap. Moreover we can observe that out two MC methods gave larger confidence intervals than Test 2-Static: probably, the parameters used for Test 2B-Static shape a harder pricing problem than previous test, and more simulations should be performed to obtain same quality results. Also in this case, HPDE proved to be the most stable method.

5.1.4 Test 3-Static: Hedging

To reduce financial risks, insurance companies have to hedge the sold VA: to accomplish this target they must calculate the greeks of products.

In this test we want to show how the different methods can be used to calculate the main greeks. This can be done through finite differences for small shocks on the variable. In general, the PDE methods are ahead w.r.t. MC methods: the price is computed through finite differences and so the price under shock is already computed. For MC methods this is quite harder because the pricing has to be repeated changing the inputs.

To start, we calculate the underlying greek delta, for the products of Test 1-Static and Test 2-Static. As in this case we don't want to compute the fair fee α_g , we fix it arbitrarily. We choose two values for each model: one for no ratchet case, and one for ratchet case. The values chosen are such as to cover the costs of the insurer regardless of the correlation, and may be plausible on a real case. Results are available in Table 10 (all values in Table must be multiplied by 10^{-4}).

In this Test, we got very accurate results with all method. Anyway, HPDE proved to be the best: it is the more stable and accurate. We remark that despite fair fee changes a lot when changing the correlation parameter ρ , the value of Delta changes much less. Delta calculation proved to be harder in the Heston model case than in the BS HW model case.

5.1.5 Test 4-Static: Risk Management

Mortality and longevity risks are unhedgeable risks. Usually the Risk Management Team has to calculate the financial reserve taking into account these risks. Usually extreme scenarios are chosen and policies are priced according to them. In this test we analyzed how the different pricing methods behaved under mortality shocks: the mortality probabilities have been increased by 10% except the last one who's equal to 1. To be brief, we simply report the fair fee for D case. Results are available in Table 11.

In this Test, we got results similar to Test 1-Static and Test 2-Static, and mortality shocks didn't affect the convergence quality of the four methods. We observe that mortality shocks reduce the value of α_g (about minus 5 bp) and this means that an increase in mortality shouldn't be a source of losses for the insurer. Consequently, insurers should pay attention to longevity risk.

$\overline{S_0}$	r	curve	k	ω	ρ	σ
100	0.04	flat	1.0	0.2	variable	0.15

Table 4: The model parameters about Test 1-Static.

			no ra	tchet				annual 1	ratchet		
ρ		HMC	$_{\mathrm{SMC}}$	$_{ m HPDE}$	APDE	ВМ	HMC	SMC	$_{ m HPDE}$	APDE	ВМ
	A	45.99 ± 1.06	45.58 ± 0.73	45.72	47.35	45.81	84.85 ± 1.23	84.58 ± 0.85	84.56	88.20	84.71
0.5	B	45.52 ± 0.79	45.31 ± 0.38	45.71	46.09	40.61	84.35 ± 0.91	84.15 ± 0.44	84.60	85.66	04.71
1	C	45.58 ± 0.52	45.71 ± 0.17	45.69	45.99	± 0.12	84.31 ± 0.60	84.65 ± 0.20	84.63	85.36	± 0.14
	D	45.85 ± 0.38	45.71 ± 0.08	45.72	45.81	±0.12	84.68 ± 0.44	84.63 ± 0.10	84.64	84.94	±0.14
	A	82.29 ± 1.39	81.40 ± 0.95	81.87	82.91	81.88	157.77±1.65	156.77 ± 1.14	156.36	161.04	157.09
0	B	82.43 ± 1.04	81.53 ± 0.50	81.92	81.75	01.00	157.68±1.23	156.55 ± 0.59	156.46	157.91	137.09
	C	81.62 ± 0.68	81.77 ± 0.22	81.80	81.89	10.16	156.50±0.80	157.05 ± 0.27	156.87	157.70	10.10
	D	81.99 ± 0.50	81.83 ± 0.11	81.79	81.81	± 0.16	157.16±0.59	157.07 ± 0.13	156.96	157.27	±0.19
	A	111.75±1.76	110.30±1.20	111.14	109.23	111.05	224.19±2.15	222.14±1.48	221.78	227.14	222.83
+0.5	B	112.73 ± 1.32	110.85 ± 0.63	111.07	108.93	111.05	224.59 ± 1.60	$222.26{\pm}0.77$	222.32	223.44	222.83
Ŧ	C	110.89 ± 0.86	111.08 ± 0.28	111.05	109.93	10.20	222.18±1.05	222.94 ± 0.35	222.52	223.36	10.24
	D	111.29 ± 0.63	111.11 ± 0.14	111.02	110.42	±0.20	222.97±0.77	$222.94{\pm}0.17$	222.67	222.96	± 0.24

	HMC	SMC	HPDE	APDE
A	30 s	30 s	30 s	28 s
B	119 s	$120 \mathrm{\ s}$	$128 \mathrm{\ s}$	$184 \mathrm{\ s}$
C	$472 \mathrm{\ s}$	$478 \mathrm{\ s}$	395 s	$461~\mathrm{s}$
D	$1866 \mathrm{\ s}$	$1896~\mathrm{s}$	$1903~\mathrm{s}$	$1800~\mathrm{s}$

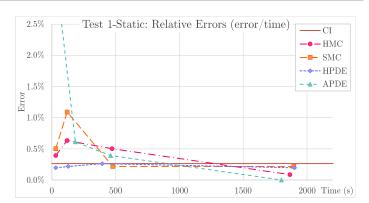


Table 5: Test 1-Static. In the first Table, the fair fee for the Black-Scholes Hull-White model, with no ratchet or annual ratchet. In the second Table the run times for the no-ratchet case ($\rho=-0.5$). Finally, the plot of relative error (w.r.t. MC value) for the three methods in the case $\rho=-0.5$ with no ratchet. The parameters used for this test are available in Table 2 and in Table 4.

	V_0				ρ	r
100	0.15^{2}	0.15^{2}	1.0	0.2	variable	0.04

Table 6: The model parameters about Test 2-Static.

			no ra	tchet				annual	ratchet		
ρ		HMC	SMC	HPDE	APDE	BM	HMC	SMC	HPDE	APDE	BM
	A	36.77 ± 1.25	36.17 ± 1.36	37.00	37.41	37.16	61.47 ± 1.23	60.90 ± 1.35	61.51	62.30	61.84
0.5	B	36.74 ± 0.92	36.40 ± 0.99	37.01	37.26	37.10	61.29 ± 0.90	60.85 ± 0.97	61.59	62.06	01.84
ı	C	36.79 ± 0.59	36.94 ± 0.62	37.01	37.11	±0.12	61.36 ± 0.58	61.56 ± 0.61	61.63	61.80	± 0.11
	D	37.47 ± 0.43	37.33 ± 0.42	37.01	37.06	±0.12	62.15 ± 0.42	61.95 ± 0.42	61.66	61.77	±0.11
	A	35.67 ± 1.56	34.02 ± 1.61	35.18	35.52	35.22	63.22 ± 1.60	61.63 ± 1.65	62.56	63.43	62.64
0	B	34.53 ± 1.13	34.48 ± 1.22	35.18	35.39	39.22	61.88 ± 1.17	61.64 ± 1.25	62.55	63.11	02.04
	C	35.05 ± 0.74	35.05 ± 0.77	35.15	35.24	10.15	62.39 ± 0.76	62.37 ± 0.79	62.59	62.78	10.11
	D	35.28 ± 0.54	35.47 ± 0.53	35.15	35.19	± 0.15	62.47 ± 0.54	62.97 ± 0.55	62.59	62.68	± 0.11
	A	33.70 ± 2.02	32.43 ± 2.14	32.58	32.76	32.63	61.44 ± 2.06	63.26 ± 2.31	62.84	63.99	62.97
+0.5	B	31.45 ± 1.43	32.26 ± 1.64	32.58	32.77	32.03	62.58 ± 1.61	62.41 ± 1.72	62.90	63.62	02.97
Ŧ	C	32.63 ± 0.96	32.94 ± 0.99	32.54	32.68	10.10	63.53 ± 1.02	62.94 ± 1.06	62.88	63.22	10.20
	D	32.31 ± 0.69	33.00 ± 0.72	32.52	32.65	±0.19	62.55 ± 0.83	62.43 ± 0.86	62.89	63.09	± 0.20

	HMC	SMC	HPDE	APDE
A	30 s	30 s	32 s	30 s
B	$122 \mathrm{\ s}$	$119 \mathrm{\ s}$	$131 \mathrm{\ s}$	114 s
C	$477 \mathrm{\ s}$	$476 \mathrm{\ s}$	$410 \mathrm{\ s}$	$491 \mathrm{\ s}$
D	$1915 \mathrm{\ s}$	$1907~\mathrm{s}$	$1755~\mathrm{s}$	$1933~\mathrm{s}$



Table 7: Test 2-Static.In the first Table, the fair fee for the Heston model, with no ratchet or annual ratchet. In the second Table the run times for the no-ratchet case ($\rho=-0.5$). Finally, the plot of relative error (w.r.t. MC value) for the three methods in the case $\rho=-0.5$ with no ratchet. The parameters used for this test are available in Table 2 and in Table 6.

		θ		ω	ρ	r
100	0.22^{2}	0.22^{2}	4.75	0.55	-0.569	0.04

Table 8: The model parameters about Test 2B-Static.

		No rat	chet			$Annual\ ratchet$				
	HMC	SMC	HPDE	APDE	BM	HMC	SMC	HPDE	APDE	BM
A	138.54 ± 2.70	141.86 ± 2.70	130.83	137.13	131.11	125.40 ± 2.46	128.33 ± 2.77	117.08	124.19	117.56
B	132.57 ± 1.98	137.16 ± 2.18	130.80	135.76	131.11	119.33 ± 1.80	123.68 ± 2.00	117.18	122.78	117.30
C	131.10 ± 1.28	135.79 ± 1.32	130.80	133.85	±0.80	117.74 ± 1.17	124.49 ± 1.20	117.23	120.09	+0.71
D	130.22 ± 0.92	132.17 ± 0.90	130.82	133.02	±0.80	116.98 ± 0.84	118.72 ± 0.82	117.19	119.62	±0.71

Table 9: Test 2B-Static. Fair fee for the Heston model, with no ratchet or annual ratchet. The parameters used for this test are available in Table 3 and in Table 8.

				no ratche	$t \ (\alpha_g = 1)$	50)			ratchet ($\alpha_g = 250$))	
	ρ		HMC	SMC	HPDE	APDE	MC	HMC	SMC	HPDE	APDE	MC
		A	6055 ± 12	6060 ± 8	6058	6034	6059	7123 ± 12	7119 ± 8	7118	7128	7121
]jj	0.5	B	6043 ± 9	6054 ± 4	6058	6050	0059	7108 ± 9	7115 ± 4	7119	7126	(121
		C	6059 ± 6	6057 ± 2	6058	6052		7117 ± 6	7119 ± 2	7120	7120	1.1
Hull-White		D	6059 ± 4	6057 ± 1	6058	6055	±1	7120 ± 4	7119 ± 1	7120	7120	±1
		A	6057 ± 13	6057 ± 9	6060	6026	6059	7390 ± 13	7392 ± 9	7380	7394	7393
lole		B	6052 ± 10	6057 ± 5	6059	6044	0059	7382 ± 10	7389 ± 5	7387	7391	1393
Scb	0	C	6058 ± 6	6057 ± 2	6058	6050	1.1	7389 ± 6	7392 ± 2	7390	7389	±1
Black-Scholes		D	6059 ± 5	6058 ± 1	6058	6055	±1	7391 ± 5	7392 ± 1	7391	7390	±1
Bla		A	6095 ± 13	6093 ± 9	6100	6002	6097	7647 ± 14	7651 ± 9	7636	7649	7650
	0.5	B	6101 ± 10	6097 ± 5	6098	6041	6097	7646 ± 10	7648 ± 5	7643	7644	7000
	+	C	6098 ± 7	6096 ± 2	6097	6063		7647 ± 7	7651 ± 2	7647	7644	1.0
		D	6099 ± 5	6097 ± 1	6097	6080	±1	7650 ± 5	7650 ± 1	7649	7646	±2

				no ratch	et $(\alpha_g = 5)$	50)			ratchet ($\alpha_g = 100$)	
	ρ		HMC	SMC	HPDE	APDE	MC	HMC	SMC	HPDE	APDE	MC
		A	7870 ± 20	7856 ± 23	7875	7867	7875	8509 ± 14	8499 ± 15	8502	8512	8509
	0.5	B	7873 ± 15	7868 ± 16	7875	7873	1813	8506 ± 10	8503 ± 11	8506	8516	8909
	1	C	7874 ± 9	7877 ± 10	7875	7874	±1	$.8505 \pm 7$	8511 ± 7	8507	8512	±1
		D	7888 ± 7	7880 ± 7	7875	7872	Ξ1	8513 ± 5	8513 ± 5	8508	8506	Τ1
Heston		A	7803 ± 23	7181 ± 25	7797	7786	7897	8405 ± 16	8390 ± 17	8395	8400	8398
He		B	7792 ± 16	7790 ± 18	7797	7794	1691	8398 ± 12	8391 ± 13	8397	8405	0990
	0	C	7796 ± 11	7801 ± 11	7797	7797	± 2	8399 ± 8	8398 ± 8	8397	8401	± 2
		D	7803 ± 8	7803 ± 8	7797	7795	Ξ2	8402 ± 6	8403 ± 6	8398	8395	Ξ2
		A	7730 ± 31	7719 ± 31	7718	7699	7718	8268 ± 22	8292 ± 22	8281	8283	8282
	+0.5	B	7703 ± 20	7717 ± 22	7717	7712	1110	8292 ± 16	8281 ± 16	8282	8290	0202
	+	C	7718 ± 13	7726 ± 14	7717	7717	±3	8283 ± 10	8484 ± 10	8282	8286	± 2
		D	7714 ± 9	7723 ± 11	7717	7715	±3	8278 ± 7	8287 ± 7	8282	8279	±2

Table 10: Test 3-Static. Delta calculation for the Black-Scholes Hull-White model and the Heston model, with no ratchet or annual ratchet (these value must be multiplied by 10^{-4}). The parameters used for this test are available in Table 2, in Table 4 and in Table 6.

				No ratche	et				Ratchet		
		HMC	SMC	HPDE	APDE	BM	HMC	SMC	HPDE	APDE	BM
	-0.5	41.89	41.75	41.75	41.85	41.84	75.96	75.91	75.92	76.20	76.00
>	-0.5	± 0.37	± 0.08			± 0.12	± 0.42	± 0.09			± 0.13
ΜH		76.73	76.21	76.18	76.22	76.28	143.12	143.03	142.95	143.23	143.06
BS	0	±0.48	± 0.11			± 0.15	± 0.56	± 0.12			± 0.18
	105	104.65	104.48	104.42	103.89	104.41	204.49	204.45	204.24	204.49	204.37
	+0.5	±0.61	± 0.13			± 0.19	± 0.73	± 0.16			± 0.23

				No ratch	et				Ratchet		
		HMC	SMC	HPDE	APDE	BM	HMC	SMC	HPDE	APDE	BM
	-0.5	34.48	34.19	33.88	33.92	33.88	56.03	56.83	55.54	55.61	55.56
_		± 0.43	± 0.42			± 0.11	± 0.42	± 0.41			± 0.11
Heston	0	32.03	32.23	31.90	31.94	31.88	55.89	55.18	55.80	55.87	55.85
Hes		± 0.53	± 0.53			± 0.14	± 0.55	± 0.54			± 0.15
	+0.5	28.97	29.67	29.25	29.32	29.26	55.11	55.93	55.40	55.53	55.39
		± 0.68	± 0.71			± 0.19	± 0.72	± 0.72			± 0.20

Table 11: Test 4-Static. Impact of +10% mortality shocks of fair fee. The parameters used for this test are available in Table 2, in Table 4 and in Table 6.

5.2 Dynamic case

In the Dynamic case, the policy holder is supposed to choose the worst strategy from an hedger point of view, changing the value of γ_t . The PH can withdraw more $(1 \le \gamma_t \le 2)$ or less $(0 \le \gamma_t \le 1)$ than the standard rate (see 2.4.3 for more details).

In this pricing framework, we refer to the prices in Forsyth and Vetzal [9]: in their article, the authors price a GLWB contract in a static framework, under the Black Scholes model with r=0.04 and $\sigma=0.15$. The contract parameters are reported in the Table 12 (Table 6.7 in [9]). They treated two cases: no ratchet, and ratchet every 3 years; both of them corresponds to case 4 in Section 4.1.3 and 4.3.4. In the first case they got $\alpha_g=63.1$ bp and in the second case $\alpha_g=70.7$ bp. In Test 1 and 2 we introduce respectively stochastic interest rate and stochastic volatility to analyze the impact of these model developments on the fair guarantee fee. The parameters for interest rate and volatility models are the same as the static case.

Here a brief summary of the numerical results for this Section.

5.2.1 Test 1-Dynamic: the Black-Scholes Hull-White Model

Test 1-Dynamic is the dynamic case of Test 1-Static. Model parameters are shown in Table 4. Results are available in Table 13.

In this Test, PDE methods proved to be more efficient than MC ones. In fact MC ones use Longstaff-Schwartz method to find the optimal withdrawal: this method needs a lot of scenarios to approximate through the least squares approach the value of the police for a given set of variable, and the regression is time demanding. Then, working at fixed time, we could perform fewer scenarios than static case (around 10%), while PDE methods used almost the same parameters as in static case. Moreover the regression problem proved to be be hard: sometimes, excluding the value $\gamma = 0$ among the possible values that the PH can chose (therefore excluding no withdrawal case), we got higher values for α_q . This means that the regression isn't very accurate, and sometimes we fail to find the optimal withdrawal: that's why, using MC methods we usually find smaller value for α_g than the right value. In particular, we excluded the value $\gamma = 0$ while using configurations A and B. We would remark that also the benchmark is affected by these computation problems and in case no-ratchet with $\rho = -0.5$ we got a small value for benchmark then PDE method (around 261 vs 266). Another thing to remark is that MC methods behaved better while ratchets were considered: maybe in this case in it easier to find the best strategy. The two MC methods proved to be equivalent: the differences in scenario generation's run-time are negligible because most of the time is spent in finding the best withdrawal. Both APDE and HPDE method gave good and stable results, but HPDE performed better in case A.

5.2.2 Test 2-Dynamic: the Heston Model

Test 2-Dynamic is the dynamic case of Test 2-Static. Model parameters are shown in Table 6. Results are available in Table 14.

In this Test, things are similar to Test 1-Dynamic, but the optimization problem seemed to be easier than in Test 1-Dynamic: MC methods converged better, especially when using high level configurations. PDE methods behaved good as usual, and HPDE method proved to be a bit better then APDE method. The

Initial age of PH	65	Gr. premium	100	DB payment	$\operatorname{cont.ly}$	Strategy	Dynamic	
G	0.05	Initial fees		Ratchet	Off/On	Bonus	5%	
Withdrawal rate	1 per Y	α_m	0	Ratchet rate	every 3 Ys	$\kappa(t)$	see tab below	
First withdrawal	1^{st} anniv.	Fees taken	cont.ly					
	$\left \begin{array}{c}\kappa\left(t\right)\end{array}\right \begin{array}{c}0\leq t\\5$	$\begin{array}{c c} t \le 1 & 1 < t \le 2 \\ \% & 4\% \end{array}$	$ \begin{vmatrix} 2 < t \le \\ 3\% \end{vmatrix}$	$3 \mid 3 < t \le 4$ 2%	$\left \begin{array}{c} 4 < t \le 5 \\ 1\% \end{array} \right $	t > 5 $0%$		

Table 12: The contract parameters used in the dynamic case.

two MC methods proved to be equivalent. We note that, in Heston model case, Dynamic strategy increase the value of α_g less than in BS HW case: probably, playing on interest rate, let the PH to gain more than playing on volatility.

5.2.3 Test 3-Dynamic: Hedging

Test 3-Dynamic is the dynamic case of Test 3-Static. Results are available in Table 15.

In this Test, we got good results with all methods, but MC methods proved to be inaccurate while using configurations A and B. The range of possible values for Delta increased with regard to Test 3-Static. The two MC methods proved to be equivalent.

5.2.4 Test 4-Dynamic: Risk Management

Test 4-Dynamic is the dynamic case of Test 4-Static. Results are available in Table 16.

In this Test, we got similar results with regard to Test 4-Static: the fees reduced a little (around 20 bp in the BS HW model case and around 6 bp in the Heston model case).

In Figure 5.1, we present, as an example, the optimal strategy at time t = 1 in two different cases. We can see the best strategy at time t = 1. We can see how it is worth to lapse when the account value reaches high values, and especially when interest rate is high or volatility low. It's more difficult to understand when do no withdrawal: there must be a convenient mix of all the variables.

			no rat	chet				ratch	net		
ρ		HMC	SMC	HPDE	APDE	ВМ	HMC	SMC	HPDE	APDE	ВМ
	Α	79.71 ± 9.68	66.5 ± 10.1	85.50	81.92	84.35	98.5 ± 16.1	90.24 ± 10.5	96.57	92.80	96.37
0.5	B	77.21 ± 5.10	72.78 ± 4.52	85.34	84.78	64.55	92.31 ± 5.12	92.69 ± 5.01	96.34	95.75	90.57
1	C	79.86 ± 2.73	80.29 ± 2.60	85.27	84.86	±0.66	94.70 ± 2.77	93.79 ± 2.67	96.25	95.78	±0.64
	D	81.66 ± 1.58	81.58 ± 1.46	85.23	84.54	±0.00	93.66 ± 1.62	94.78 ± 1.62	96.19	95.40	±0.04
	A	162.6 ± 18.5	148.4 ± 13.2	172.55	167.86	169.05	182.3 ± 13.2	179.5 ± 14.3	186.44	181.96	186.53
0	B	155.43 ± 7.70	156.52 ± 6.95	172.60	171.48	109.05	182.42 ± 6.16	181.35 ± 6.53	186.48	185.33	180.55
	C	161.53 ± 4.39	164.88 ± 4.41	172.57	171.61	±0.90	184.21 ± 3.70	183.84 ± 3.58	186.54	185.45	±0.86
	D	164.51 ± 2.53	163.15 ± 2.23	172.58	171.15	±0.90	183.87 ± 2.21	183.27 ± 2.15	186.55	185.00	±0.80
	Α	256.6 ± 15.6	238.6 ± 23.3	265.33	261.22	261.29	273.9 ± 24.8	262.7 ± 24.9	272.18	268.10	274.02
+0.5	B	246.0 ± 10.4	248.5 ± 10.9	266.66	265.71	201.29	268.37 ± 8.39	269.82 ± 9.11	273.24	272.33	274.02
Ť	C	253.70 ± 5.96	253.33 ± 5.38	266.93	265.94	11.02	271.90 ± 5.51	271.60 ± 4.49	273.67	272.46	11.20
	D	259.00 ± 3.38	254.06 ± 3.11	267.29	265.38	±1.23	272.24 ± 2.95	270.35 ± 2.86	273.99	271.94	±1.20

	HMC	SMC	HPDE	APDE
A	30 s	$31 \mathrm{\ s}$	32 s	30 s
B	$119 \mathrm{\ s}$	$122 \mathrm{s}$	$127 \mathrm{\ s}$	$120 \mathrm{\ s}$
C	$482 \mathrm{\ s}$	$487 \mathrm{\ s}$	$463 \mathrm{\ s}$	466 s
D	$1911 \mathrm{\ s}$	$1942 \mathrm{\ s}$	$1732~\mathrm{s}$	$1815 \mathrm{\ s}$

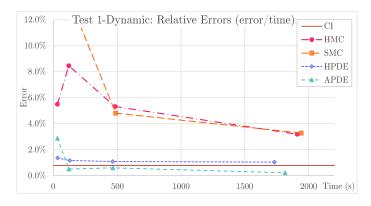


Table 13: Test 1-Dynamic. In the first Table, the fair fee for the Black-Scholes Hull-White model, with no ratchet or annual ratchet. In the second Table the run times for the no-ratchet case ($\rho=-0.5$). Finally, the plot of relative error (w.r.t. MC value) for the three methods in the case $\rho=-0.5$ with no ratchet. The parameters used for this test are available in Table 12 and in Table 4.

			no rat	\overline{chet}				rate	het		
ρ		HMC	SMC	HPDE	APDE	BM	HMC	SMC	HPDE	APDE	BM
	A	56.96 ± 6.01	62.56 ± 8.78	64.57	64.89	65.38	66.72 ± 8.37	72.06 ± 6.17	71.23	71.58	72.03
0.5	B	61.68 ± 3.05	57.20 ± 3.96	64.72	64.64	05.38	72.53 ± 3.56	68.45 ± 3.36	71.37	71.30	72.03
Ī	C	64.68 ± 2.02	64.03 ± 2.05	64.76	64.42	± 0.45	71.65 ± 1.92	71.95 ± 1.99	71.43	71.04	± 0.45
	D	63.85 ± 1.28	64.67 ± 1.28	64.81	64.35	±0.45	70.39 ± 1.22	71.33 ± 1.25	71.50	70.96	±0.45
	A	58.83 ± 11.0	58.68 ± 19.68	61.92	61.92	62.32	65.95 ± 9.04	79.17 ± 10.36	68.97	68.98	69.54
	B	58.95 ± 4.01	54.34 ± 4.08	61.91	61.67	02.32	69.14 ± 4.53	65.30 ± 4.25	68.95	68.69	09.54
0	C	58.26 ± 2.25	57.76 ± 2.34	61.88	61.43	±0.56	68.89 ± 2.43	68.50 ± 2.44	68.94	68.41	±0.58
	D	59.16 ± 1.43	59.70 ± 1.47	61.87	61.35	±0.50	66.76 ± 1.53	68.67 ± 1.57	68.93	68.33	±0.58
	A	52.66 ± 15.29	61.70 ± 13.04	57.25	57.50	50.50	63.67 ± 9.76	83.75 ± 13.01	64.60	64.79	CT 40
+0.5	B	57.26 ± 5.06	51.95 ± 6.09	57.33	57.26	56.59	68.21 ± 6.30	60.31 ± 5.94	64.66	64.53	65.42
Ŧ	C	52.23 ± 3.01	51.47 ± 3.50	57.36	57.01	10.72	63.78 ± 3.13	64.06 ± 3.25	64.68	64.25	10.74
	D	52.60 ± 1.85	52.24 ± 1.83	57.39	56.94	± 0.73	62.98 ± 1.87	61.86 ± 1.97	64.71	64.16	± 0.74

	HMC	SMC	HPDE	APDE
A	30 s	31 s	33 s	28 s
B	$119 \mathrm{\ s}$	$122 \mathrm{\ s}$	126 s	$107 \mathrm{\ s}$
C	481 s	493 s	$418 \mathrm{\ s}$	$460 \mathrm{\ s}$
D	$1903 \mathrm{\ s}$	$1844~\mathrm{s}$	$1690~\mathrm{s}$	$1896 \mathrm{\ s}$

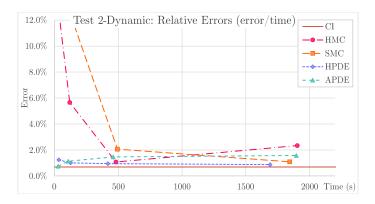


Table 14: Test 2-Dynamic. In the first Table, the fair fee for the Heston model, with no ratchet or annual ratchet. In the second Table the run times for the no-ratchet case ($\rho=-0.5$). Finally, the plot of relative error (w.r.t. MC value) for the three methods in the case $\rho=-0.5$ with no ratchet. The parameters used for this test are available in Table 12 and in Table 6.

				no ratchet	$\alpha_g = 30$	00)			ratchet (e	$\alpha_g = 350$)	
	ρ		HMC	SMC	HPDE	APDE	BM	HMC	SMC	HPDE	APDE	BM
يو		A	8513 ± 573	7951 ± 557	8078	8097	8078	8959 ± 542	8279 ± 551	8148	8200	8157
ji	0.5	B	8220 ± 347	8304 ± 305	8081	8091	0070	7922 ± 324	8104 ± 330	8151	8180	8197
	Ī	C	8091 ± 172	8146 ± 161	8082	8073	±42	8155 ± 173	8120 ± 162	8152	8163	±38
Hull-White		D	8089 ± 108	8105 ± 104	8082	8093	T42	8164 ± 101	8174 ± 99	8152	8164	±36
		A	7898 ± 550	7697 ± 489	7516	7531	7485	7733 ± 509	7379 ± 667	7538	7539	7517
lole		B	7685 ± 257	7389 ± 269	7517	7529	1460	7640 ± 259	7417 ± 249	7527	7539	1911
Sch	0	C	7488 ± 139	7443 ± 150	7517	7514	±31	7433 ± 145	7604 ± 137	7528	7530	±29
Black-Scholes		D	7428 ± 90	7489 ± 83	7517	7518	±31	7523 ± 87	7525 ± 84	7528	7533	±29
Bla		A	7444 ± 470	7612 ± 491	7333	7342	7224	7569 ± 421	7292 ± 500	7304	7314	7309
_	0.5	B	7257 ± 242	7368 ± 209	7337	7350	7324	7386 ± 207	7413 ± 192	7308	7322	7309
	+0	C	7306 ± 116	7201 ± 124	7339	7336	±28	7469 ± 116	7293 ± 112	7309	7307	±26
		D	7270 ± 78	7302 ± 78	7340	7339	± 26	7291 ± 70	7298 ± 68	7310	7310	±20

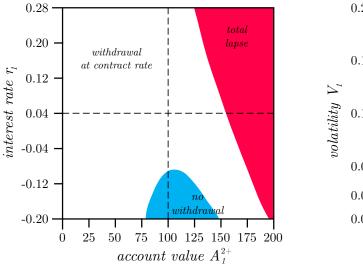
				$no\ ratchet$	$(\alpha_g = 75)$	5)			ratchet (e	$\alpha_g = 100$)	
	ρ		HMC	$_{\mathrm{SMC}}$	$_{ m HPDE}$	APDE	$_{\mathrm{BM}}$	HMC	$_{\mathrm{SMC}}$	$_{ m HPDE}$	APDE	ВМ
		A	8181 ± 472	7794 ± 524	8436	8429	8432	8349 ± 293	8374 ± 324	8477	8474	8481
	0.5	B	8440 ± 250	8383 ± 223	8436	8436	0432	8304 ± 174	8527 ± 154	8479	8480	0401
	1	C	8405 ± 87	8426 ± 87	8437	8437	1.10	8535 ± 80	8499 ± 76	8479	8480	±14
		D	8437 ± 50	8472 ± 53	8437	8438	±19	8516 ± 44	8501 ± 43	8480	8480	±14
Heston		A	8756 ± 626	8304 ± 586	8329	8319	8297	8751 ± 758	8440 ± 420	8341	8332	8351
He		B	8080 ± 283	8345 ± 394	8330	8327	0291	8466 ± 251	8184 ± 217	8341	8339	0331
		C	8313 ± 148	8137 ± 185	8330	8329	±29	8303 ± 106	8398 ± 101	8341	8340	±18
		D	8330 ± 75	8238 ± 79	8330	8331	±29	8343 ± 52	8283 ± 66	8341	8441	±18
		A	7308 ± 1145	7453 ± 914	8217	8205	8242	8244 ± 736	8192 ± 522	8191	8180	8206
	0.5	B	8238 ± 623	7919 ± 416	8218	8215	0242	8150 ± 276	8213 ± 229	8191	8189	8200
	+	C	8143 ± 145	7874 ± 241	8218	8217	±46	8216 ± 133	8123 ± 178	8192	8190	±22
		D	8144 ± 119	8131 ± 108	8218	8219	±40	8123 ± 66	8195 ± 77	8192	8191	122

Table 15: Test 3-Dynamic. Delta calculation for the Black-Scholes Hull-White model and the Heston model, with no ratchet or ratchet (once every 3 years); these value must be multiplied by 10^{-4} . The parameters used for this test are available in Table 12, in Table 4 and in Table 6.

				No ratche	et				Ratchet		
		HMC	SMC	HPDE	APDE	BM	HMC	SMC	HPDE	APDE	BM
	-0.5	76.64	75.74	76.85	76.19	76.16	85.10	85.47	86.59	85.84	86.97
	-0.5	± 1.49	± 1.37			± 0.60	± 1.53	± 1.51			± 0.61
HW		149.80	149.34	155.47	154.11	155.97	168.14	167.94	169.78	168.26	170.32
BS	0	±2.38	± 2.14			± 0.96	± 2.14	± 2.00			± 0.96
		236.49	236.33	242.08	240.22	242.1	248.40	246.49	250.08	248.12	250.24
	+0.5	±2.84	± 2.79			±1.2	± 2.68	± 2.48			± 0.53

		No ratchet					Ratchet				
		HMC	SMC	HPDE	APDE	BM	HMC	$_{\mathrm{SMC}}$	HPDE	APDE	$_{\mathrm{BM}}$
Heston	-0.5	57.74	58.77	58.78	58.32	58.56	63.73	64.50	64.74	64.22	64.59
		±1.26	± 1.27			± 0.46	± 1.23	± 1.27			± 0.46
	0	53.08	53.77	55.57	55.07	55.76	59.86	61.17	61.79	61.22	62.11
		± 1.45	± 1.46			± 0.58	± 1.51	± 1.59			± 0.59
	+0.5	46.33	46.19	50.93	50.50	50.66	55.67	54.79	57.29	56.76	59.01
		±1.87	± 1.77			± 0.70	± 1.85	± 1.94			± 0.76

Table 16: Test 4-Dynamic. Impact of +10% mortality shocks of fair fee. The parameters used for this test are available in Table 12, in Table 4 and in Table 6.



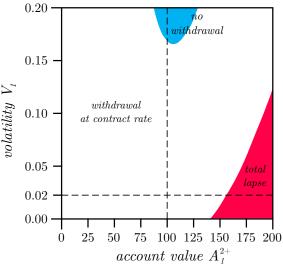


Figure 5.1: Optimal strategy at the first event time (t=1) for the BS HW model and the Heston model, assuming $B_1^{2+}=100$. Model parameters are available in Tables 4 and 6. Product parameters are available in Table 12, and $\alpha_g=135$ bp for both cases.

6 Conclusions

In this article we have developed four methods to price GLWB contracts under different conditions. Regarding the stochastic model, both stochastic interest rates and stochastic volatility effects have been considered. Regarding the policy holder behavior, both static and dynamic strategy have been considered.

Since GLWB variable annuities are such a long maturity products, the effects of stochastic interest rates and stochastic volatility cannot be overlook. In particular, the impact of stochastic rate seems to be more relevant. Also Forsyth and Vetzal in [9] used regime switching model having both stochastic interest rate and volatility, but our approach, based on SDE, is more realistic, and suitable for hedging.

All four methods gave compatible results both for pricing and delta calculation. The fair hedging fee (i.e. the cost of maintaining the replicating portfolio) is determined using a sequence of parameters' refinements. The PDE methods proved to be not very expensive, while MC methods proved to be more expensive. The Hybrid PDE seemed to be the more performing than the others for its convergence speed and stability of results. Also ADI PDE behaved very well but the implementation was harder then Hybrid PDE one. In the BS HW model case, Standard MC thanks to its exact simulation outperformed the Hybrid Method while, in the Heston model case, the MC methods proved to be roughly equivalent, even if the Hybrid MC was easier to be implemented.

As we said before, PDE methods proved to be much more efficient than MC methods, especially in Dynamic case where is much more simple to implement the optimal withdrawal choice. Similarity reduction reduces the problem dimension to a 2D problem and therefore PDE methods perform well. Anyway, we have to remark that MC methods offer a confidence interval for the result, they are useful in risk measures calculation (for example VAR or ES), and they are preferred by insurance companies because of their attachment to the idea of scenario.

A future development that could be treated is to combine stochastic interest rate and stochastic volatility: the combined model could be an element of greater realism.

We conclude by pointing out that our methods are quite flexible in that they can accommodate a wide variety of policy holder withdrawal strategies such as ones derived from utility-based models.

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